Inflation risk premia in the US and the euro area

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Motivation

- Nominal and real bond yields are often used by central banks to extract information of relevance for monetary policy.
- Break-even inflation rates used as an indicator of markets' inflation expectations / a measure of central bank credibility
- Break-even rates are a noisy measure due to the existence of premia; notably: inflation risk premia (but also zero-coupon effects, liquidity premia)
- How large are inflation premia in the US and the euro area? What are their properties/determinants?

Outline of the presentation

- Methodology
- Data and estimation
- Results: inflation risk premia; raw vs. premium-adjusted break-even rates; properties and macro determinants

Inflation risk premia: what can we expect?

- Inflation risk premium filtered from the overall term premium filtered from yields: plenty of measurement uncertainty
- Guidance from intuition/theory? Inflation risk premium should be proportional to inflation variability; positive and increasing in maturity
 - in theory a negative premium is possible; sign depends on covariance between inflation (real returns on bonds) and SDF (consumption)
 - recent empirical results point to mostly positive inflation premia, but the magnitude varies: + and sizeable in Buraschi and Jiltsov (2005) and Ang et al. (2008); small but + for long maturities in Durham (2006) and D'Amico et al. (2008). Euro evidence in Garcia and Werner (2008): small and +.

Methodology

- Important to include relevant information: index-linked and macro data; survey data
- Consistency with macro dynamics: inflation expectations derived from a "structural," empirically plausible macro-model
- No-arbitrage restrictions added to the macro structure: the nominal and real term structures are modelled to be arbitrage-free and mutually consistent

 \implies Macro-based interpretations of term structure dynamics, including dynamic responses of yields, premia, and break-even inflation rates to structural shocks

The model: Macro

inflation:
$$\pi_t = \bar{\pi} + \mu_\pi E_t \pi_{t+} + (1 - \mu_\pi) \delta_\pi \pi_{t-} + \delta_x x_t + \varepsilon_t^\pi$$

output gap:
$$x_t = \mu_x E_t x_{t+} + (1 - \mu_x) \zeta_x x_{t-} - \zeta_r (r_t - E_t [\pi_{t+1}]) + \varepsilon_t^x$$

short rate:
$$r_t = \bar{r} + (1 - \rho) \left[\beta \left(E_t \pi_{t+12} - \pi_t^* \right) + \gamma x_t \right] + \rho r_{t-1} + \eta_t$$

infl. target: $\pi_t^* = \phi_{\pi^*} \pi_{t-1}^* + u_t^{\pi^*}$

Key ingredients of the model

Macro: REE solution

State variables $(\pi, x, r, \pi^*, ...)$ $\mathbf{Z}_t = \mathbf{M}\mathbf{Z}_{t-1} + \boldsymbol{\xi}_t$ Policy rate $r_t = \boldsymbol{\delta}\mathbf{Z}_t$

Finance: assumption on stochastic discount factor

 $\begin{array}{ll} \text{Market prices of risk} & \lambda_t = \lambda_0 + \lambda_1 \mathbf{Z}_t \\ \implies \text{Nominal yields} & \mathbf{Y}_{n,t} = \mathbf{A}_n + \mathbf{B}_n \mathbf{Z}_t \\ \implies \text{Real yields} & \mathbf{Y}_{n,t}^* = \mathbf{A}_n^* + \mathbf{B}_n^* \mathbf{Z}_t \end{array}$

Estimation

- Bayesian Maximum Likelihood using Kalman Filter; we exploit prior information on structural economic relationships
- Real yields enter the likelihood function late in the sample (US 2003; euro area 2004) to reduce initial liquidity problems
- Survey data information (inflation and short-term interest rate) explicitly included in the estimation
- Estimation using simulated annealing to reduce risk of local maxima

Data - US sample: January 1990 — April 2008

- Macro data: y-o-y inflation, output gap (log-GDP in deviations from CBO estimate of potential; ARIMA forecast/interpolation)
- Nominal yields: 1-, 3-, 6-, 12-m, 3-, 5-, 10-y zero-coupon rates (Fed Board)
- Real yields: 3-, 5-, 7-, and 10-y zero-coupon rates extracted from US TIPS (as of 2003)
- Survey data: 3-m rate and inflation in 2/4 quarters and next 10 years (SPF)

Data - euro sample: January 1999 — April 2008

- Macro data: y-o-y inflation, output gap (log-GDP in deviations from a quadratic trend, as in CGG98; ARIMA forecast/interpolation)
- Nominal yields: 1-, 3-, 6-, 12-m, 3-, 5-, 10-y zero-coupon rates extracted from German bond prices and EUR money market rates
- Real yields: 3-, 5-, 7-, and 10-y zero-coupon rates extracted from French and German HICP-linked bonds
- Survey data: 3-m rate in 3/12 months (Consensus); inflation 1, 2, 5 years ahead (SPF)





Inflation and estimated inflation target: euro area



Term structure of average risk premia: US



Term structure of average risk premia: euro area



Estimated 10-year US premia



Estimated 10-year euro area premia







Euro area 10-year break-even inflation rates and survey inflation forecasts



Estimated US 10-year inflation risk premium and output gap



Estimated US 10-year inflation risk premium and inflation



Estimated euro 10-year inflation risk premium and output gap



Estimated euro 10-year inflation risk premium and inflation











Conclusions

- Using a macro-finance model for real and nominal term structure dynamics, we provide estimates of the size and dynamics of inflation risk premia in the US and the euro area
- Our framework allows us to obtain macro interpretations of term structure and premia dynamics
- Inflation risk premia are positive and relatively small on average; they vary over time, mostly in response to output gap and inflation changes
- Break-even inflation rates are therefore a noisy measure of inflation expectations, but, inparticular in the US, much of their variation seem to be due to changes in expected inflation

Extra slides

Estimated 10-year US premia



Estimated 10-year euro area premia







The inflation risk premium

For the short-rate, we can write $r_{t} = r_{t}^{*} + E_{t} [\pi_{t+1}] + prem_{\pi,t}^{1} + \frac{1}{2} \Sigma_{\pi} \Sigma_{\pi}'$ where $prem_{\pi,t}^{1} = -\Sigma_{\pi} (\lambda_{0} + \lambda_{1} \mathbf{X}_{1,t}) \quad \text{inflation premium}$

 $\Sigma_{\pi} \equiv C_{\pi}\Sigma$ "amount of risk" $\lambda_t = \lambda_0 + \lambda_1 X_{1,t}$ "price of risk" The model: Market prices of risk

Market prices of risk determined empirically as affine functions of the states

$$oldsymbol{\lambda}_t = oldsymbol{\lambda}_0 + oldsymbol{\lambda}_1 imes States_t$$

We use

$$\boldsymbol{\lambda}_{t} = \begin{pmatrix} \lambda_{01} \\ \lambda_{02} \\ \lambda_{03} \\ \lambda_{04} \end{pmatrix} + \begin{pmatrix} \lambda_{11} & \lambda_{12} & \lambda_{13} & \lambda_{14} \\ \lambda_{21} & \lambda_{22} & \lambda_{23} & \lambda_{24} \\ \lambda_{31} & \lambda_{32} & \lambda_{33} & \lambda_{34} \\ \lambda_{41} & \lambda_{42} & \lambda_{43} & \lambda_{44} \end{pmatrix} \begin{pmatrix} \pi_{t}^{*} \\ r_{t} \\ \pi_{t} \\ x_{t} \end{pmatrix}$$

Each row of λ_t determines the price of risk associated with each of the states; these vary over time with the level of the states

Figure 1a: US nominal zero-coupon yields



Figure 1b: Euro area nominal zero-coupon yields



Figure 2a: US real zero-coupon yields











Figure 3b: Euro area zero-coupon break-even inflation rates

