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Abstract

Little is known about the extent and drivers of information flow within couples and whether spouses hold aligned expectations about the same outcomes. To provide new evidence, we conduct an online survey of 2,200 middle-aged married couples in the U.S. Our focus is on expectations about Social Security benefits. We first document misalignment in expectations: the correlation between partners' beliefs about a given spouse's Social Security benefits is 0.70, well below full agreement. We also show that this imperfect alignment is systematically associated with couple-specific characteristics. To establish causal evidence on information spillovers, we implement a randomized information experiment paired with a sequential survey design, where the index spouse receives targeted information, and the other is surveyed a few days later. Our findings reveal that information provided to the index spouse partially spills over to their partner, with the average treatment effect on the second spouse's expectations being about half that observed for the index spouse. Using detailed survey data on measures of communication frictions, cognitive barriers, and the value of information, we identify key drivers of information flow. Spillovers are larger when communication barriers are low and when the information is particularly valuable. We also show that the information treatment enhances conditions for better intra-household decision-making. Overall, our findings highlight the importance of incorporating realistic communication frictions into models of household decision-making.

JEL classification: D83, D84, D13, J12

Key words: information spillovers, expectations, household decisions, frictions, Social Security benefits, retirement

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1 Introduction

Many economic decisions – how much to save, when to retire, how much labor to supply – are made jointly at the household level. The traditional workhorse model of household decision-making has been the unitary model, which assumes a single utility function for the entire household (Samuelson, 1956; Becker, 1974, 1981). In recent decades, economists have shifted from the unitary model to the collective approach in modeling household decision-making; these models recognize that the household is composed of individuals with different preferences. An important assumption in the collective model is complete information, meaning that information flows freely between partners, leading them to hold the same expectations about future outcomes (e.g., Lundberg and Pollak 1996; Chiappori and Mazzocco 2017). Despite the centrality of this assumption, empirical evidence on the alignment of spousal expectations for the same outcomes and the extent of information flow within couples remains limited. This remains the case even as expectations data have become increasingly available in surveys (Bachmann et al., 2023).

In this paper, we aim to fill this knowledge gap by conducting an online survey of 2,200 middle-aged married couples in the US, in which both spouses report their expectations about the same outcomes. We have three primary research objectives. First, we use these novel data to test the assumption of equality of expectations within couples. Second, we quantify the extent of information spillovers between partners and investigate the factors that facilitate or hinder information flow. Third, we explore how the provision of information facilitates decision-making within couples. Investigating information transmission within couples in real-life settings is challenging. To address this, we implement a randomized information intervention alongside a sequential survey design. Specifically, one treated spouse receives decision-relevant information exogenously, with no explicit incentive to share information. The second spouse is invited to complete the survey only after their partner has responded, allowing information diffusion to occur endogenously as it does in the “wild.” By comparing the expectations of the second spouse in treated couples (that is, couples where the first spouse is treated with information) with those in control couples (where no decision-relevant information is provided), we identify the extent of spillovers.

Our primary focus is on Social Security benefits expectations. This is motivated by several factors. First, Social Security benefits are a crucial component of financial security in retirement in the United States, with half of those aged 65 and older receiving at least 50% of their income from Social Security benefits, and a quarter relying on it for at least 90% (Dushi et al., 2017). Second, there is widespread evidence of misinformed and uncertain expectations regarding individual Social Security benefits, as well as pensions in other contexts (e.g., Mitchell 1988; Guiso et al. 2013; Liebman and Luttmer 2015; Caplin et al. 2022). Third, expectations and uncertainty about individual Social Security benefits have been shown to play an important role in retirement planning (e.g., Bottazzi et al. 2006; Delavande and Rohwedder 2011; Luttmer and Samwick 2018).¹ Fourth, de-

¹Similarly, Social Security policy features or reforms, which change people’s expectations about future Social Security benefits, have been shown to play an important role in retirement decision-making (e.g., Liebman et al., 2009; Mastrobuoni, 2011; Coyne et al., 2024; Deshpande et al., 2024).

cisions regarding savings and retirement timing are often made jointly by spouses (e.g., Blundell et al., 2016). Moreover, couples may adopt strategies to maximize lifetime Social Security wealth, so decision-making depends not only on an individual's beliefs about their own benefits but also on their partner's expectations.

We start out with a descriptive analysis of expectations within the control group. First, we show that individual expectations deviate significantly from the estimates provided by the Social Security Quick Calculator (SS Quick Calculator) based on earnings. On average, respondents expect to receive \$2,350 per month if they claim at age 67, while they anticipate their spouse's age 67 benefits to be \$2,270. For both their own and their spouse's benefits, approximately 45% of respondents overestimate their expected benefits compared to the Quick Calculator estimate, while about 50% underestimate them. The remaining 5% report expectations within \$50 of the calculator's projection, indicating that only a small fraction have expectations closely aligned with the calculator forecast.

Next, we show significant within-couple disagreement about a given spouse's benefit. Specifically, when comparing spouses' reported expectations for the same benefit, only 35% of couples have a difference of less than \$100, while nearly 50% have a discrepancy greater than \$200. The correlation of expectations for a given spouse's benefit is 0.70. We can rule out measurement error as the sole driver of this imperfect alignment. Consistent with communication frictions playing some role in this imperfect correlation, we see that the correlation is (weakly) lower in couples with lower levels of marital satisfaction, with below-median marital length, and with greater disagreement about the quality of the marital relationship (which is presumably indicative of some dysfunction in the relationship). This imperfect correlation is not due to Social Security benefits being unimportant to the household's finances: couples for whom benefits are likely to be a major source of post-retirement income (i.e., couples with below-median earnings) do not exhibit stronger correlations in their expectations. Moreover, we confirm that this level of correlation is not exclusive to Social Security benefit expectations. Our findings reveal similar correlations ranging from 0.46 to 0.64 for other decision-relevant expectations, including unemployment, inflation, stock market prices, household spending growth, and earnings growth.²

We then provide causal evidence on information spillovers within a couple by leveraging a randomized information intervention combined with the sequential survey design within couples. The information treatment involves providing the index spouse (the first spouse we interview) with a forecast for future Social Security benefits for themselves and their spouse, calculated using a method similar to the SS Quick Calculator. The treatment leads the index spouses to revise their expectations (about own and spouse's Social Security benefits) substantially, reducing the absolute gap between expectations and the forecast by an average of 22 percentage points (39% of the

²An imperfect correlation by itself does not necessarily imply inefficiency. Couples, for example, may specialize where one spouse knows the benefits and makes financial decisions, and the other does not (rationally) invest in the acquisition of Social Security information. We can rule this out in our data, since we find that spouses tend to be similarly misinformed. That is, if one partner overestimates or underestimates benefits, the other is also likely to do the same.

baseline gap). Subsequently, we investigate whether this information spills over to the secondary spouse. Consistent with some information diffusion, having a treated spouse leads to a secondary spouse – who, on average, takes the survey a couple of days after the index spouse – having an absolute perception gap that is 10-12 percentage points lower compared to those in the control group. Thus, almost half of the treatment effect on the index spouse spills over to the secondary spouse, a pattern observed for both the index and secondary spouse’s benefits.

Importantly, we investigate the heterogeneity in information spillovers to secondary spouses to understand the factors that facilitate or hinder information flow. Standard economic theory suggests that valuable information is more likely to be shared, so we focus on indicators of informational value, such as perceived credibility and baseline accuracy. We also explore strategic motivations, particularly whether news that affects bargaining power is more or less likely to be shared. Given the complexity of the Social Security system, we also investigate cognitive frictions, which may be more prevalent among couples with lower financial literacy or education.³ Beyond these, some couples may face higher communication costs, either due to poor relationship quality, diverging interests or recurrent conflicts. As proxy for these costs, we examine disagreement on other dimensions, such as marital quality and household decision-making roles. Couples who align on these aspects are likely to communicate more effectively, suggesting lower communication costs. Finally, we assess how couple-level characteristics (e.g., gender of the index spouse, marital length) shape information flow dynamics.

We find significant heterogeneity in information spillovers across key dimensions. Spillovers are larger in couples where the information is more valuable. Additionally, spillovers increase when the index spouse learns that their own or their partner’s benefits are lower than expected. This finding contradicts strategic motives for concealing unfavorable information, which would predict that individuals might withhold bad news about themselves to maintain bargaining power. Instead, for many couples, the cost of not adjusting financial plans in response to lower-than-expected benefits may be high, incentivizing greater information sharing.⁴ Importantly, spillovers are significantly smaller in couples with markers of communication frictions. Finally, we find no role for cognitive frictions. Overall, our findings support a cooperative household model with information frictions, where sharing is more likely when the potential gains are present and when communication barriers are minimal.

Last, we investigate whether the information treatment enhances conditions for better intra-household decision-making. For this purpose, we focus on the alignment in expected claiming ages for one spouse (that is, the distance between the age at which spouse A expects to claim and the age at which spouse B expects spouse A to claim). It is unlikely that a couple is optimizing

³Choukhmane et al. (2023) find that 20% of married couples in their sample could improve outcomes by shifting Defined Contribution plan contributions from the lower to the higher marginal match incentive spouse. Their survey evidence suggests half of these misallocations are accidental, highlighting cognitive constraints in joint retirement planning.

⁴This could also be consistent with negative news being more salient, though evidence on asymmetric updating in response to good and bad news is quite mixed (Benjamin, 2019).

their financial planning for retirement if one expects their spouse to retire and claim much later or earlier than what is intended by the spouse themselves. We find that this alignment of expected claiming age increases in the treatment group for couples whose expectations were misaligned at the baseline. This suggests that providing information to one spouse—and the resulting spillovers to the other—can create conditions more conducive to optimal intra-household decision-making.

An important contribution of this paper is to test the often-implicit assumption of alignment of economically-relevant expectations within couples. In the last few decades, there has been growing research on individual decision-making under uncertainty using subjective expectations data elicited in surveys (Manski, 2004). This body of work has shown that expectations play a crucial role in individual decision-making (see Bachmann et al. 2023 for a recent literature review). However, due to data limitations, few studies have compared expectations about the same outcomes from spouses. Our paper is the first to document this correlation in expectations about individual-level outcomes and behaviors that are central for household economic decision-making. Our findings challenge the standard assumption of full information within couples and underscore the importance of incorporating information frictions into theoretical models. They also highlight the need to measure within-household heterogeneity in expectations when modeling economic behavior and designing interventions.

We also know very little about how expectations within the same couple or decision-making unit influence choices. Few studies have explored this dynamic. Notable exceptions include Giustinelli (2016), which examines teenagers’ and parents’ school choice under subjective risk about future choice-specific outcomes; Almås et al. (2024) and Walker et al. (2024), which discuss intra-household allocation of parental investments under uncertainty about the productivity of those investments; and Ke (2025), which examines how disagreement in macroeconomic expectations affects household portfolio choice. Beyond expectations, imperfect knowledge about spousal preferences may also play a key role in joint decision-making (Michaud et al., 2020). Our study contributes to this work by demonstrating how information can impact alignment in claiming age plans, and helps identify factors that govern efficient information flows, both in terms of household characteristics and potential interventions.

Our work also builds on a growing literature using within-survey randomized information provision, which has shown that survey respondents are not always well-informed about important variables and that they are responsive to new information (see Fuster and Zafar, 2023; Haaland et al., 2023 and Stantcheva, 2023 for recent reviews).⁵ This literature has primarily focused on how information provided to one person influences the beliefs or decisions of that person. In contrast, we study how information flows beyond the survey to other household members. This approach provides new insights into the process of expectations formation in real life—an area where significant gaps in understanding remain (Manski, 2023)—and, in particular, highlights the

⁵Existing studies have focused on a wide range of topics, including contraception (Delavande, 2008), education (Wiswall and Zafar, 2015), inflation (Coibion et al., 2020), discrimination (Haaland and Roth, 2023), and tax policy (Stantcheva, 2021).

role of the household as an important but imperfect source of information sharing.

Finally, our study also contributes to an emerging literature on information transmission within couples through exogenously-provided information (Conlon et al. 2022; Apedo-Amah et al. 2023; Ashraf et al. 2023; Fehr et al. 2023). Focusing on a context where spouses may find it beneficial to share, Conlon et al. (2022) study a stylized lab-in-the-field setting in India, and Fehr et al. (2023) focus on learning about household income rank in Germany in a survey. Both studies focus on gender dynamics and reveal that husbands' information is more effectively transmitted than wives'. Our findings suggest that this gender disparity may vary depending on the context and outcomes under consideration.⁶ Specifically, we find only weak evidence of gender-based heterogeneity in information spillover for a key economic outcome crucial for many households. Spillovers are slightly larger when the index spouse is male, but these differences are not statistically significant.⁷ We are not aware of any work to date in developed countries' settings that empirically investigates how information relevant for key economic decisions (such as saving, retirement, labor supply) diffuses within households. In addition, our rich data allow us to identify factors that impede or facilitate the flow of information within couples when there are gains to sharing. Prior work has solely focused on the gender dynamics of information-sharing.

2 Data and Context

2.1 Social Security in the U.S.

The Old-Age, Survivors, and Disability Insurance (OASDI) program provides financial assistance to eligible individuals, focusing primarily on retirees, disabled individuals, and survivors. The program pays monthly retirement benefits to eligible workers. Eligibility for benefits is primarily contingent upon an individual's work history and contributions to Social Security through payroll taxes. Benefits are typically computed using average indexed monthly earnings. This average summarizes up to 35 years of a worker's indexed earnings. Eligible individuals can start claiming benefits as early as age 62, but this may lead to a reduction of up to 30% in monthly benefits compared to claiming at the *full retirement age* (67 for those born after 1960). Claiming benefits after the full retirement age may result in larger monthly benefits, with maximum benefits attained at age 70 (Social Security Administration, 2020).

⁶Consistent with context being relevant, Mustafi (2024) finds that, in a stylized lab-in-the-field setting in India, households pool information successfully when the information pertains to a gender-specific domain, but not when it is gender-neutral.

⁷Our prior is that spouses have an incentive to share information about Social Security benefits, and we will test this by considering strategic motives. Apedo-Amah et al. (2023) and Ashraf et al. (2023) focus on information flow in situations where spouses may have individual strategic incentives to withhold information because of differing preferences, e.g. regarding fertility. This work builds on research from development economics documenting barriers to spousal communication when spouses have private information they can strategically use to influence household decisions. This includes studies of temporary migrants who can easily hide pertinent information like earnings (e.g., Chen 2013; de Laat 2014; Ur Rehman 2023) or studies where the observability of choices or resources is experimentally manipulated (e.g., Ashraf et al. 2014).

A married person aged 62 or older can qualify for benefits based on both their own work record and their spouse's work record, given a minimum one-year marriage. Spousal benefits range from one-third to one-half of the worker's benefits. If the individual is also eligible for benefits based on own earnings, they will receive the higher benefit amount between own and spousal benefits (Social Security Administration, 2021). Divorced individuals who were married at least 10 years are eligible for Social Security benefits based on their ex-spouse's record if they are unmarried and not entitled to a higher benefit on their own record when they become eligible.⁸

Various claiming options therefore exist for married couples, and determining the optimal strategy is complex, depending on factors such as each spouse's life expectancy, earnings history, health, discount rate, and leisure preferences. An essential aspect of formulating this strategy, as underscored by numerous online retirement planning advisors, involves a full understanding of both one's future Social Security benefits and those of one's spouse.

2.2 Research Design and Sample

Our survey was administered by NielsenIQ between January and July 2022. We recruited respondents from two separate pools used by NielsenIQ: rotated-out married respondents from the NY Fed's Survey of Consumer Expectations who previously reported they would consent to NielsenIQ contacting their spouses for future surveys, and a third-party panel of respondents NielsenIQ uses for other research projects. In recruiting these individuals whom we refer to as "index spouses" henceforth, we restricted the available sample to those who are between ages 38 and 62, not retired, not receiving any income from Social Security, and reporting that their spouses are also not receiving any income from Social Security. Lastly, we only included index spouses who shared an email address that is different than their own for their spouses.

After index spouses were screened, they received the link to the survey with an incentive of \$15. In the introduction of the survey, respondents were informed that their spouses ("secondary spouses", henceforth) will also receive a link to a similar survey with an incentive of \$15 and that if their spouse completes the survey, the couple (specifically, the index spouse) will receive an additional \$15. The secondary spouse survey links were automatically sent once the index spouses completed their survey. This sequential design allows us to measure spillovers within a household. The average gap between the index spouse completing the survey and the secondary spouse starting the survey is 2.5 days in our sample.

In the survey, both index and secondary spouses received randomized information treatments with slightly different designs, as described in Section 2.3. We assigned 22% of index spouses to the control group and 78% to the treatment group. A much larger proportion of the sample was assigned to the treatment group since, within the treatment group, secondary spouses are randomized into two different kinds of treatments. In addition, since much of our analysis exploits

⁸In the event of a spouse's or ex-spouse's death, the surviving widow is entitled to receive widow's benefits. The widow is eligible to receive a percentage ranging from 71 percent (at age 60) to the full 100 percent (at full retirement age) of the amount the deceased spouse was receiving before their passing (Social Security Administration, 2021).

within-individual variation, the treatment group share was chosen to be larger. The treatment assignment is the same for both members of the couple, so that if the index spouse is in the control group, the secondary spouse is also assigned to the control group.

This recruiting process led to 9,159 index spouse and 2,219 couple completes (that is, about a quarter of the index spouses had their partner also complete the survey). In our empirical analysis, we use the sample of complete couples. Secondary spouses in treatment couples are slightly more likely to complete the survey than those in control couples, with a completion rate of 25.5% versus 22.6%. This difference is statistically significant (p -value = 0.004). In Appendix Table A1, we regress the likelihood of the secondary spouse completing the survey onto a rich set of characteristics of the index spouse, separately for the control and treatment couples. The selection (on observables) for secondary spouse completion is similar for the two groups (none of the p -values testing the equality of coefficients between treatment and control groups in the last column of the table are below 0.10). This suggests little evidence of differential selection on observables between treatment and control couples.

Additionally, Table A1 shows that, in both the treatment and control groups, secondary spouses are generally more likely to complete the survey in couples where the index spouse is more financially literate and has smaller absolute perception gaps in own expected benefits (as we discuss later, this is defined as the gap between the respondent's expected benefits and benefits as reported by the Quick Calculator). Importantly, marital satisfaction, as reported by the index spouse, does not significantly influence the likelihood of a couple completing the survey. This indicates that our sample is not biased toward couples with particularly high levels of marital satisfaction.

The demographic characteristics of our sample of couples is presented in Table 1. The majority of the index spouses are female and the majority of the sample is white, has a college degree and is below age 45. Over three-quarters of spouses work full-time. On average, couples have been married for 17 years, and one-third currently live with children under 18. Due to its importance in retirement planning, the survey elicited measures of financial literacy. Approximately 30% of the sample is considered financially literate, based on their ability to correctly answer two questions regarding compound interest and the real value of money under inflation.⁹ We also measure perceived financial savviness by asking, "How financially savvy would you say you are?" On average, respondents rate themselves as 5 on a 7-point scale.

Related to marital satisfaction, we elicit respondents' perceptions about their relationship with their spouses using four questions on how happy they are with their relationship, whether they have a warm and comfortable relationship with their spouse, how rewarding they find their relationship with their spouse, and how satisfied they are with their relationship; the responses to

⁹The first question asks "Let's say you have \$200 in a savings account. The account earns 10% interest per year. Interest accrues at each anniversary of the account. If you never withdraw money or interest payments, how much will you have in the account at the end of two years?". The second question asks "Imagine that the interest rate on your savings account was 1% per year and inflation was 2% per year. After one year, how much would you be able to buy with the money in this account?". These questions have been widely used in the prior literature on financial literacy (Lusardi and Mitchell 2014).

these are elicited with a 5 or 6-point Likert scale. These questions are from the 4-item battery developed by Funk and Rogge (2007). Using the responses to these questions, we create a marital satisfaction index defined as the sum of the standardized values for these questions (the index is also standardized to have mean 0 and variance 1, with a higher value indicating greater satisfaction). There is substantial variation across individuals and within couples in the index (Appendix Table A2). For example, of the index spouses whose responses put them in the highest quartile of the index, only about two-thirds of their spouses also report marital satisfaction in the same quartile. In fact, as shown in the table, 43% of couples answer such that both spouses fall in *different* quartiles of the marital satisfaction index. We leverage this within-couple variation later as a proxy for disagreement and communication frictions within the couple. There are limited studies looking at disagreement on marital satisfaction within couples. However, such disagreement has been found to correlate with a higher likelihood of divorce (Gager and Sanchez, 2003).

TABLE 1: Demographic Characteristics of the Sample

| | Index Spouse | | | Secondary Spouse | | |
|---------------------------------------|--------------|---------|---------|------------------|---------|---------|
| | Control | Treated | P-Value | Control | Treated | P-Value |
| Female | 0.59 | 0.58 | 0.92 | 0.41 | 0.43 | 0.51 |
| White | 0.88 | 0.89 | 0.66 | 0.86 | 0.88 | 0.24 |
| Age < 45 | 0.55 | 0.53 | 0.51 | 0.53 | 0.49 | 0.21 |
| Age 45–50 | 0.21 | 0.19 | 0.37 | 0.21 | 0.21 | 0.88 |
| Age > 50 | 0.24 | 0.28 | 0.11 | 0.27 | 0.30 | 0.21 |
| College degree | 0.62 | 0.66 | 0.10 | 0.57 | 0.61 | 0.13 |
| % Financially literate | 0.29 | 0.32 | 0.24 | 0.29 | 0.32 | 0.18 |
| % Working full-time | 0.77 | 0.75 | 0.18 | 0.80 | 0.76 | 0.18 |
| % With kids | 0.31 | 0.35 | 0.13 | 0.31 | 0.35 | 0.13 |
| Length of marriage | 16.77 | 17.35 | 0.19 | 16.77 | 17.35 | 0.19 |
| Perception of own financial savviness | 5.15 | 5.32 | 0.04 | 4.97 | 5.07 | 0.24 |
| Marital Satisfaction Index | −0.10 | −0.04 | 0.22 | 0.01 | 0.07 | 0.26 |
| Observations | 492 | 1,727 | | 492 | 1,727 | |

Note: Financial literacy is defined based on having the two financial questions included in the survey correct. There are four questions that ask participants to self-evaluate their relationship with the spouse in the survey. Marital satisfaction is an index created by the sum of the standardized values for these 4 questions. Financial savviness is measured with a 7-point Likert scale.

Table 1 also shows that the control and treatment groups are balanced in terms of background characteristics for both index and secondary spouses. The only statistically significant difference is that index spouses in the treatment group report slightly higher average financial savviness, though the difference is economically small.

Table A3 shows the comparison of the demographic characteristics of the index spouses in our sample to the married, non-retired household heads in the Current Population Survey (CPS). The sample of index spouses is broadly representative of the married, non-retired household heads in the U.S. based on geography and employment status. The index spouses in our sample, on average, are younger, more educated and are more likely to be female—these are all patterns that tend to be observed in online samples. When we compare the weekly earnings of the index spouses in our sample to the household heads in the CPS, we observe that college-graduate index spouses earn slightly more than their CPS counterparts while those without a college degree earn slightly less than the household heads in the CPS. One potential reason for the discrepancies we note be-

tween the two samples is the additional restriction of not earning any Social Security income in our sample. Unfortunately, we do not have information on this for the respondents in the CPS.

2.3 Survey Design

We describe the Social Security expectations module and information treatment here. The questionnaire and some screenshots can be found in Appendix B.

The **index spouse's** module started with a series of introductory questions regarding the percent chance of both themselves and their spouse being eligible to receive Social Security benefits, and from which record these benefits would be claimed. The module proceeded in the following stages (for a visual depiction, see Figure B1):

Stage 0: Index spouse baseline Social Security expectations. Respondents answered several Social Security-related expectations. These included their own and their spouse's expected monthly benefits, conditional on claiming at ages 62, 67, and 70 (e.g., *If you were to retire at age 67 and collect benefits, on average, how much do you expect the monthly payments to be in today's dollars?*; see Figures B2 and B3). Respondents also reported their expected claiming ages for themselves and their spouse, as well as the expected benefit amounts conditional on claiming at those ages.¹⁰

Stage 1: Information on own Social Security benefits to index spouses. Respondents in the treatment group were presented with a table showing their projected Social Security benefits if they were to claim at ages 62, 67 or 70 (see Figure B4a). These projections are derived based on the earnings they reported, utilizing a formula similar to the one used by the SS Quick Calculator.¹¹ The screen also looks similar to what individuals can see when consulting the Quick Calculator. Respondents in the control group were provided placebo information on men's soccer around the world (see Figure B5a).

Stage 2: Index spouse midline Social Security expectations. Respondents were asked their expected benefits for own and spouse conditional on claiming at their (potentially updated) expected claiming age, and conditional on claiming at age 67 (if different from expected claiming age).

Stage 3: Information on spouse's Social Security benefits to index spouses. Respondents in the treatment group were presented with a table showing the projected Social Security benefits for their spouse with a claiming age of 62, 67, and 70 (see Figure B4b). These estimates were derived from the earnings reported by the respondent for their spouse, utilizing the SS Quick Calculator formula. The display featured the spouse's benefits at the top and, as a reminder, the identical table from stage 2 at the bottom of the screen (for example, see Figure B4a). Respondents in the

¹⁰For the expected claiming age and benefits, respondents were further asked a density of their beliefs. This involved indicating the likelihood of claiming within specific age ranges, or the benefits falling within designated intervals, as a way of measuring uncertainty in beliefs.

¹¹The SS Quick Calculator is available online and provides a projection of Social Security benefits after users input their date of birth and earnings in the current year: <https://www.ssa.gov/OACT/quickcalc>.

control group were provided a placebo information on swimming in the Olympic games (see Figure B5b).

Stage 4: Index spouse endline Social Security expectations. Respondents were asked the same set of questions as in stage 2.

After the index spouse completed the survey, the secondary spouse was invited to participate. The Social Security module for the **secondary spouse** commenced with identical introductory questions as those in the index spouse’s questionnaire and proceeded with the following stages:

Stage 5: Secondary Spouse baseline Social Security expectations. Respondents were asked the same set of questions as in the Index spouse Stage 0. This is Stage 5 of the experiment when viewed at the couple level but it is the first stage for the secondary spouse.

Stage 6: Information to secondary spouses. Half of the secondary spouses in the treatment group received information about projected benefits for both themselves and their spouse (T1 in Figure B1). The other half also received information about their spouse’s claiming age expectations (T2). Control participants again viewed placebo content.

Stage 7: Secondary spouse endline Social Security expectations. Respondents were asked the same set of questions as in stage 2.

Our primary interest is in understanding how information provided to the index spouse transmits to the secondary spouse. Accordingly, our main empirical analysis on spillovers focuses on data from Stages 0 to 5, but we also use data from stages 6 and 7 when analyzing alignment in intended behavior.

On average, index spouses took 38.6 minutes to complete the survey (median: 33 minutes), while the secondary spouses took 28.6 minutes (median: 23 minutes).

2.4 Summary Statistics of Social Security Expectations

Table 2 displays the summary statistics for monthly Social Security expectations reported by index spouses at baseline (Stage 0). On average, respondents expect to receive \$2,018 if claiming at age 62, \$2,346 at age 67 (Full Retirement Age), and \$2,691 at age 70. While respondents recognize that early claiming reduces benefits and delaying increases them, they underestimate the size of these adjustments on average. The expected reduction at age 62 is about 14% (versus the actual 30%), and the expected gain at age 70 is around 15% (compared to the actual 24%). Index spouses, on average, expect to claim at age 66.5, and 16.7% expect to claim at age 67. They expect their spouse to claim at age 66, on average, and the age 67 expected benefits for their spouse is

\$2,271. Full descriptive statistics for both own and spousal expectations are provided in Appendix Table A4.

TABLE 2: Index Spouse Social Security Expectations at Baseline (Stage 0)

| | Index Spouse | | | |
|---|--------------|---------|---------|---------|
| | All | Control | Treated | P-Value |
| Expected own age 62 benefits (\$) | 2017.6 | 1910.0 | 2048.2 | 0.19 |
| Expected own age 67 benefits (\$) | 2345.9 | 2227.5 | 2379.6 | 0.15 |
| Expected own age 70 benefits (\$) | 2690.9 | 2609.4 | 2714.2 | 0.37 |
| Expected SS benefit claiming age | 66.5 | 66.6 | 66.5 | 0.39 |
| Expected age 67 benefits for the spouse (\$) | 2270.7 | 2279.3 | 2268.2 | 0.91 |
| Expected SS benefit claiming age for the spouse | 66.0 | 66.2 | 65.9 | 0.14 |
| Own benefits | | | | |
| Predicted age 67 benefits by the SSA calculator (\$) | 2036.5 | 2010.0 | 2044.1 | 0.51 |
| Absolute gap in own age 67 benefits (\$) | 1209.2 | 1162.5 | 1222.6 | 0.38 |
| Share overestimating | 44% | 44% | 45% | 0.71 |
| Share underestimating | 49% | 49% | 49% | 0.80 |
| Benefits for the spouse | | | | |
| Spouse's predicted age 67 benefits by the SSA calculator (\$) | 1930.7 | 1876.3 | 1946.2 | 0.15 |
| Absolute gap in spouse's age 67 benefits (\$) | 1210.6 | 1198.0 | 1214.1 | 0.81 |
| Share overestimating | 45% | 47% | 44% | 0.24 |
| Share underestimating | 50% | 47% | 50% | 0.17 |
| Observations | 2,219 | 492 | 1,727 | |

Note: The table shows the average expected benefits and the average expected claiming age of the index spouses. In the bottom panel of the table, absolute gap is defined as the absolute value of the difference between the predicted benefits by the SS Quick calculator and the benefit expectations of the respondent. Table A4 presents the expectations for age 62 and 70 benefits along with predicted benefits and absolute gaps in addition to what is included in this table.

In the bottom panel of Table 2, we show how respondents' expectations are different from the age 67 benefit amounts predicted by the SS Quick Calculator. For both own and spouses' benefits, the absolute gap (i.e., the absolute value of the difference between the benefits predicted by the SS Quick Calculator and the respondent's expected benefit) is sizable and amounts to an average of \$1,210. Note that, the calculator might not be giving the most accurate prediction if individuals have private information about their past or future work behavior. For both their own and their spouse's benefits, around 45% of respondents overestimate their age 67 benefits compared to the projection from the calculator, while about half underestimate them. The remaining 5% of respondents report benefit expectations within \$50 of the SS calculator prediction. These patterns are similar across treatment and control groups.

In Figure A1, we examine the distribution of the perception gap between the prediction of age 67 benefits by the SS Quick Calculator and the respondents' age 67 benefit expectations (for themselves and for their spouses, separately) as a share of the prediction by the SS Quick Calculator. We see that index spouses' average perception gap in own age 67 benefits is -8.1% (i.e., an average overestimation of benefits by 8.1 percent) and the average gap in spouse's age 67 benefits is -7.3%. However, the distributions of gap for own and spouse's benefits are both highly dispersed: for example, 20% of index spouses overestimate their age 67 benefits by at least 50% and 19% of index spouses underestimate their age 67 benefits by at least 50%. Overall, these sizable gaps suggest potential lack of knowledge about the Social Security system.

Appendix Table A5 further illustrates how the absolute perception gap varies across different characteristics. There is no gender difference. Older and more financially-literate index spouses tend to have smaller absolute perception gaps. Similarly, those without a college degree or in households with below-median earnings also show smaller absolute gaps, possibly because Social Security benefits represent a more critical resource for them.

3 Correlation of Spousal Expectations: Descriptive Evidence

This section examines the alignment of spousal expectations for the same outcome within couples, focusing on expected Social Security benefits at age 67—the Full Retirement Age. To avoid potential spillover effects from the information treatment, we restrict the analysis to couples in the control group. The extent of misalignment observed in this group provides motivation for the subsequent information intervention.

3.1 Social Security Benefits Expectations

We begin by examining the histogram of the distance between spouses’ age 67 benefit expectations for the same spouse, presented in Figure 1a. Specifically, we compute the distance:

$$\text{Spouse A's exp about A's benefit} - \text{spouse B's exp of spouse A benefit.} \quad (1)$$

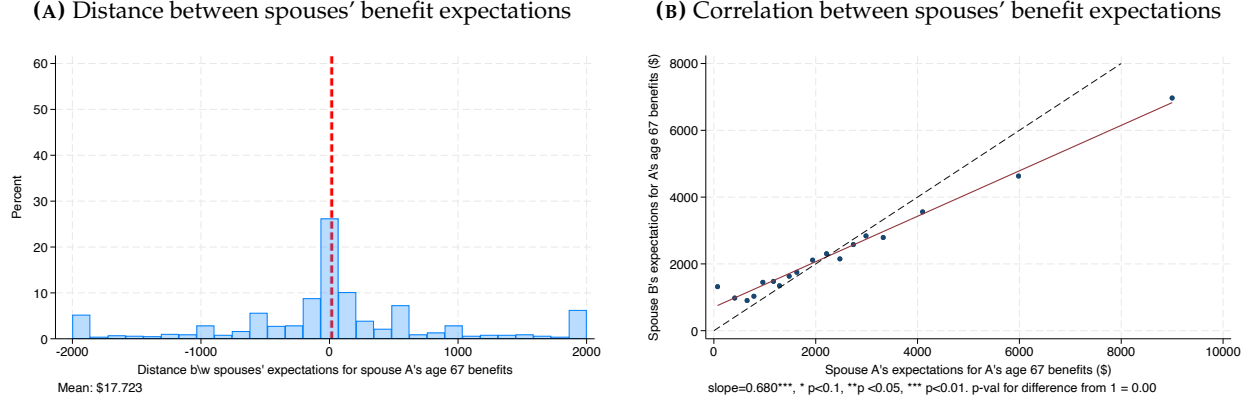
We use stage 0 expectations for the index spouse and stage 5 expectation for the secondary spouse. The distribution is nearly symmetric around zero, resulting in an average difference of approximately \$18 between spouses’ benefit expectations for the same spouse. However, the variation is substantial, with a standard deviation of \$1,540. Around 35% of respondents have distance within the $[-\$100, 100]$ range and 55% of the couples have an absolute distance of over \$200. Furthermore, the average absolute difference is \$830. This significant disparity in spouses’ expectations for the same outcome suggests that within-couple expectations may not be perfectly aligned, potentially indicating (communication or information) frictions within households.

In Figure 1b, we plot the binscatter of spouses’ expectations for the same spouse’s age 67 benefits. The correlation between spouses’ expectations is 0.70, which is statistically different than 1 and the slope of the regression line is 0.68, which is also statistically different than 1. This correlation provides further descriptive evidence against the equality of expectations within a couple.¹² When we restrict the sample to couples with benefit expectations below \$4,000,¹³ to reduce the effect of outliers on the slope of this relationship, we find that the correlation is largely unchanged at 0.66. We present a similar exercise for spouses’ expectations for Social Security wealth of the cou-

¹²This correlation is likely an upper bound since, even in the control group, taking the survey might have spurred some discussions about Social Security benefits within a couple.

¹³The maximum monthly age 67 Social Security benefits in the U.S. in 2023 were \$3,627, while the average benefit was \$1,706.

FIGURE 1: Within-Couple Expectations for Age 67 Benefits



Note: The dashed red line in the left panel shows the mean of the distance between age 67 benefit expectations for the same spouse. Both panels only use data for control group couples. For ease of exposition, in the first panel we winsorize the distance at -2000 and 2000, but report the mean of the unwinsorized distance. The right panel shows a binned scatter plot of both spouses' expectations for the same spouse's age 67 benefits. The correlation coefficient between the two variables used in the binscatter is 0.70, which is statistically different than 1 (p-value=0.00).

ple, which we define as the sum of the age 67 benefits of the spouses within a couple, in Appendix Figure A2. Here, we find the within-couple correlation in expectations to be 0.74, slightly larger than the correlation in benefit expectations. Once again this correlation is statistically different than 1, providing evidence against the equality of expectations within a couple.

But how should this correlation be interpreted—high or low? To contextualize, we compare it to a scenario where couples are randomly matched based on the observed patterns of assortative matching in our sample. Assortative matching could contribute to the observed correlation, as individuals with similar characteristics may naturally hold similar expectations, even without direct communication. When we simulate random matching of secondary spouses to primary spouses using key characteristics (gender, age, education, and employment status), we find a correlation of just 0.12 in expectations (Appendix Figure A3). This stark contrast suggests that the observed correlation is far higher than what assortative matching alone would predict, indicating that some degree of information flow does occur within households.

Even in the absence of any frictions, there are several reasons why the correlation in spouses' expectations may be imperfect. One possible explanation is that respondents report their beliefs with measurement error, which would naturally reduce the correlation to below 1. To assess whether this is the case, we perform an exercise where we assume both spouses have the same expectations, but that the second spouse reports beliefs with a classical measurement error (distributed normally with mean zero and standard deviation σ , added to log beliefs about benefits) while the index spouse reports beliefs without any errors. In this setup, the σ that would generate the correlation of 0.70 we observe in the data is 0.61, which implies that 51% of the observed variation in secondary spouse benefits is due to noise.¹⁴ To put these numbers in context, we check

¹⁴We simulate secondary spouse benefits, B_s^{me} in this setup with measurement error as $\log(B_s^{me}) = \log(B_i) + \varepsilon$, where B_i is the benefit expectations of the index spouse and $\varepsilon \sim N(0, \sigma^2)$. We first find the σ value that generates

the test-retest correlation in expectations of the control group index spouses between Stage 0 and Stage 4. We find a correlation of 0.88, which would imply a σ of 0.35 if we followed the same steps. This σ that generates the test-retest correlation implies that only 23.3% of the variation in the control group index spouse beliefs about benefits is due to measurement error.¹⁵ Given the difference in the variation in beliefs due to noise in these two cases, we conclude that measurement error is unlikely to be the sole driver of this imperfect alignment.

Lastly, these misperceptions persist even among couples who accurately perceive their spouse's earnings, a key determinant of future Social Security benefits. Among the control group, the correlation between spouses' beliefs about one spouse's earnings is 0.81. When we restrict our sample to couples with accurate earnings perceptions (i.e., those couples where there is no discrepancy between what an individual says they earn and what their spouse reports their earnings to be, representing 46.5% of our sample), the correlation in Social Security benefits expectations remains at 0.75. Again, this correlation is still statistically different than 1 (and from 0.88, which is the test-retest correlation in the control group), further indicating misalignment in expectations even among couples with accurate earnings perceptions (see Appendix Figure A4).

Another mechanism that would generate imperfect correlation in spouses' expectations is specialization within the household. If one spouse makes all the financial and economic decisions of the household, it would be natural for spouses to have different expectations for the same outcomes.¹⁶ However, our analysis of within-couple correlations of Social Security wealth, presented as a heat map in Figure 2, suggests a different story. Here we define Social Security wealth as the sum of two spouses' monthly benefits. Specifically, we assign each spouse a decile based on the size of their perception gap regarding Social Security wealth (i.e., the gap between their expected Social Security wealth and the Social Security wealth predicted by the calculator). The figure reveals that most couples cluster along the 45-degree line, indicating a high degree of correlation in their perception gaps. This finding implies that specialization within the household is unlikely to explain the observed discrepancies. In addition, the heat map suggests that "errors" might be amplified within couples, making some couples particularly vulnerable by having both members over or under-estimating their Social Security wealth.

3.2 Heterogeneity

Having established that the imperfect correlation is unlikely due to measurement error or specialization-based explanations, we next conduct heterogeneity analysis, to provide some suggestive evidence on the underlying drivers of the frictions.

the observed correlation in the data as the average of 1,000 bootstrap repetitions. We then decompose the variance of second spouse benefits as $var(B_s^{me}) = \mathbb{E}[var(B_s^{me}|B_i)] + var(\mathbb{E}[B_s^{me}|B_i])$, following the law of iterated variances. The ratio $\frac{\mathbb{E}[var(B_s|B_i)]}{\mathbb{E}[var(B_s|B_i)] + var(\mathbb{E}[B_s|B_i])}$ gives us the share of the variance of B_s , not attributed to the B_i .

¹⁵It is also important to note that the observed correlation of 0.70 for spouses' Social Security benefit expectations is statistically different from the test-retest correlation of 0.88 in the control group.

¹⁶Investing in financial knowledge can be viewed as a deliberate choice, analogous to other forms of human capital investment (Lusardi et al., 2017).

FIGURE 2: Correlation in spouses' perception gaps about Social Security wealth of the couple



Note: The figure shows a heat map of both spouses' perception gaps regarding the couple's Social Security wealth.

TABLE 3: Heterogeneity in the correlation of spouses' Social Security wealth expectations

| | Corr in Exp in SS wealth | p-value of difference from 1 | p-value of differences across groups |
|---|--------------------------------|------------------------------------|--|
| Overall | 0.740*** | 0.00 | |
| Panel A | | | |
| Both spouses below age 55 | 0.737*** | 0.00 | |
| At least one spouse above age 55 | 0.723*** | 0.00 | 0.89 |
| Couple below-median earnings | 0.449*** | 0.00 | |
| Couple above-median earnings | 0.760*** | 0.00 | 0.00 |
| At least one spouse no college | 0.654*** | 0.00 | |
| Both spouses have college degrees | 0.726*** | 0.00 | 0.51 |
| Panel B | | | |
| At least one spouse below-median marital sat. | 0.642*** | 0.00 | |
| Both spouses above-median marital sat. | 0.836*** | 0.00 | 0.02 |
| Below median marital sat. diff. | 0.767*** | 0.00 | |
| Above median marital sat. diff. | 0.673*** | 0.00 | 0.19 |
| Couple below-median marital length | 0.675*** | 0.00 | |
| Couple above-median marital length | 0.796*** | 0.00 | 0.17 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$, show the statistical difference of the correlations from 0. The last column shows the p-value of the differences in the correlation in spouses' expectations for couple's Social Security wealth across the mutually exclusive groups, as defined in the first column.

In Panel A of Table 3, we find higher correlation in Social Security wealth expectations among couples with above-median earnings. It is possible that reduced financial stress facilitates communication, though higher-earning couples also differ in other dimensions. The correlation is also weakly higher in college-educated couples. Surprisingly, the correlation is not higher in older couples, which contrasts with the idea that older couples closer to retirement would have more

frequent discussions on the topic.

Panel B examines marital quality and duration. Couples with above-median marital satisfaction have more aligned expectations, as do those with lower disagreement about their relationship quality and longer marriages, albeit weakly. This suggests that communication frictions may play a role in the imperfect correlation within couples. In section 4, we will directly investigate the factors that lead to more information spillovers using a randomized information experiment.

3.3 Correlations in Expectations for other Decision-relevant Outcomes

We now show that the imperfect within-household correlation is not unique to Social Security benefit expectations, but extends to expectations of other decision-relevant variables as well. Our survey measures expectations about the year-ahead unemployment rate, interest rates, stock prices, inflation, and housing prices. Additionally, we measure household-level expectations such as household spending and income growth over the next 12 months. Table 4 shows that misaligned expectations are not specific to Social Security benefits and the equality of expectations is, in fact, rejected for other aggregate and household-related expectations as well. The correlations in spouses' other expectations range from 0.46 to 0.64 and that they are all statistically different than 1 (and from 0.88 which is the test-retest benchmark).

TABLE 4: Correlation in spouses' other decision-relevant expectations

| | Corr in Expectations |
|----------------------------|-------------------------|
| Unemployment rate higher | 0.546*** |
| Interest rate higher | 0.464*** |
| Stock Prices higher | 0.509*** |
| Inflation | 0.388*** |
| Household income growth | 0.639*** |
| Household spending growth | 0.533*** |
| Tax payment growth | 0.631*** |
| National home price growth | 0.587*** |

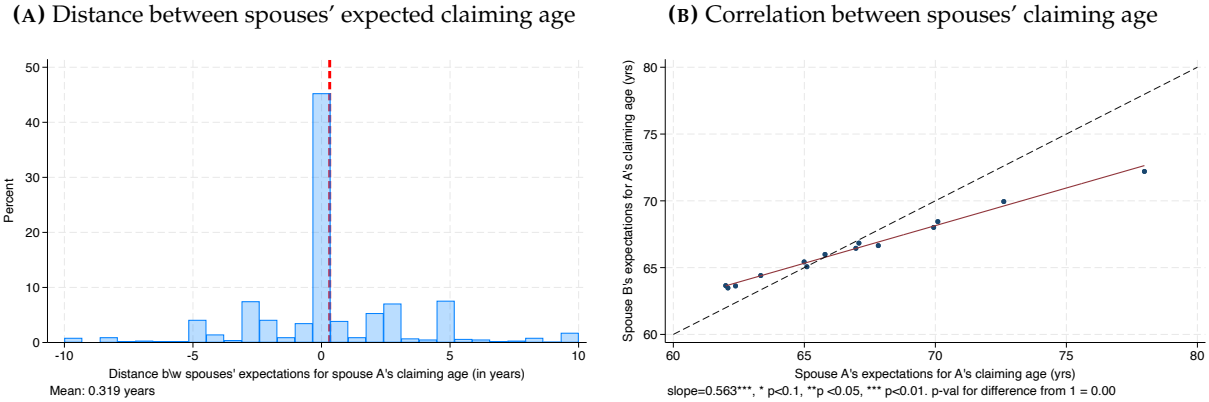
Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$, show the statistical difference of the correlations from 0. All correlations are also statistically different from 1 (p-val=0.00 for all rows). Unemployment rate higher, interest rate higher and stock prices higher denote the probabilistic expectations for the U.S. unemployment rate, average interest rate on savings account and average stock prices in the U.S. stock market being higher 12 months from now, respectively. Inflation refers to the 12-month ahead inflation expectations, national home price growth refers to the 12-month ahead expected average home price growth in the U.S. and household income and spending growth expectations refer to the 12-month ahead nominal household income and spending growth expectations. All expectations are elicited as point forecasts.

We further examine one particular type of expectation which is relevant to our study: expected Social Security claiming age. Respondents were asked at what age they expect to claim Social Security benefits, and at what age they expect their spouse to claim. The timing of claiming is an important choice variable for financial well-being in retirement, since delaying claiming increases monthly benefits up to age 70.

We depict the alignment in expected claiming age in the control group, i.e., the distance be-

tween the age at which spouse A expects to claim and the age at which spouse B expects spouse A to claim. In Figure 3, we show the distance between spouses' expected claiming age of the *same* spouse (Panel A) and the binscatter (Panel B). The left panel highlights the misalignment in expectations about claiming ages, with approximately 67% of couples having different expectations and 38% diverging by more than two years. Panel B shows that the slope of this relationship and the overall correlation are both 0.56, and again statistically different than 1.

FIGURE 3: Distance between spouses' expected claiming age



Note: The left panel shows the difference between spouses' expected claiming age for the same spouse in the control group. The dashed red line denotes the mean distance in the control group. For ease of exposition, we cap the histogram at -10 and 10 years. The right panel shows a binned scatter plot of both spouses' expectations for the same spouse's claiming age. The correlation coefficient between the two variables used in the binscatter is 0.56, which is statistically different than 1 (p-value=0.00).

4 Information Spillovers Based on a Randomized Experiment

So far, we have shown that (1) individuals are not fully informed about Social Security benefits, and (2) expectations about Social Security benefits for a given spouse are not fully aligned within couples. The descriptive patterns in section 3 suggest that some within-couple barriers may be partly responsible for this misalignment.

In this section, we investigate this directly using an information experiment. Our main goal is to see whether information randomly provided to the index spouse spills over to the secondary spouse, and to investigate the drivers of this transmission. The section concludes with a detailed heterogeneity analysis.

4.1 Conceptual Framework

We first present a simple conceptual framework to illustrate how expectations are updated within a couple following the provision of information to one spouse. For ease of exposition, we focus solely on the stage 1 information provision, and on expectations about the index spouse's Social Security benefits.

Consider the index spouse i and the secondary spouse s in couple c . The couple is assigned to either the information treatment group T or the control group C . At Stage 0, the index spouse holds prior beliefs $B_{ic(J)}^{\text{stage } 0}$, for $J = \{T, C\}$ regarding their future Social Security benefits, while the secondary spouse holds beliefs $B_{sc(J)}^{\text{stage } 0}$, for $J = \{T, C\}$ about their spouse's Social Security benefits.

At stage 1, the index spouse in the treatment group receives a signal $Signal_{ic(T)}$ about their future Social Security benefits (i.e., the predicted benefits according to the SS Quick Calculator). In a typical belief-updating framework, the posterior belief, measured at stage 2, is a weighted combination of the prior belief and the new signal:

$$B_{ic(T)}^{\text{stage } 2} = (1 - \gamma)B_{ic(T)}^{\text{stage } 0} + \gamma Signal_{ic(T)}. \quad (2)$$

The parameter γ represents the weight put on the signal, where $\gamma \in [0, 1]$. $\gamma = 0$ implies no learning, while $\gamma = 1$ implies full updating to the provided signal. Under Bayesian learning, γ is a function of the precision of the prior and the signal.

The index spouse in the control group receives a signal unrelated to their Social Security benefits so does not revise their beliefs about their future Social Security benefits:¹⁷

$$B_{ic(C)}^{\text{stage } 2} = B_{ic(C)}^{\text{stage } 0}. \quad (3)$$

To measure how close beliefs are from the signal, we define the absolute perception gap in Stage 2 as follows:

$$PG_{ic(J)}^{\text{stage } 2} = \frac{|B_{ic(J)}^{\text{stage } 2} - Signal_{ic(J)}|}{Signal_{ic(J)}} \quad J \in \{C, T\}, \quad (4)$$

where $Signal_{ic(T)}$ is the signal received by the index spouse in the treatment group. $Signal_{ic(C)}$ is the hypothetical signal index spouses in the control group would have received about their Social Security benefits had they been in the treatment group. The stage 0 perception gap is defined similarly using the prior beliefs, $B_{ic(J)}^{\text{stage } 0}$, and the information that is (or could have been) provided in stage 1.

We test whether the treatment group's Social Security benefit beliefs align more closely with the signal than the control group's by comparing the absolute perception gaps. Specifically, we exploit within-person variation between stages and focus on changes in the absolute perception gap to test the following hypothesis:

$$\textbf{Hypothesis 1: } \overline{PG_{ic(T)}^{\text{stage } 2}} - \overline{PG_{ic(T)}^{\text{stage } 0}} = \overline{PG_{ic(C)}^{\text{stage } 2}} - \overline{PG_{ic(C)}^{\text{stage } 0}},$$

where the left (right) hand side is the average revision in the treatment (control) group. In our

¹⁷To keep the exposition simple, we ignore measurement error and survey noise here. However, the empirical analysis will take them into account.

experimental design, additional information is provided to index spouses at Stage 3, and their beliefs are re-measured at Stage 4. This allows us to extend Hypothesis 1 by testing whether perception gaps further narrow between Stages 2 and 4.

While Hypothesis 1 focuses on belief revision by index spouses, our primary interest lies in how information provided to one spouse may spill over to the other. Upon receiving their signal, the index spouse i in the treatment group may choose to communicate with their spouse s , and may transmit some information $Info_{ic(T)}$. This communicated information could be either the signal she received or her updated posterior belief. We flexibly model it as follows:

$$Info_{ic(T)} = f(B_{ic(T)}^{\text{stage } 4}, Signal_{ic(T)}). \quad (5)$$

Information sharing may not be complete due to cognitive or communication frictions. The secondary spouse may then update their original prior beliefs $B_{sc(T)}^{\text{stage } 0}$ just before stage 5 (when we measure baseline beliefs for the secondary spouses). Their stage 5 beliefs can thus be written as:

$$B_{sc(T)}^{\text{stage } 5} = d_{ic(T)} \left[(1 - \delta) B_{sc(T)}^{\text{stage } 0} + \delta Info_{ic(T)} \right] + (1 - d_{ic(T)}) B_{sc(T)}^{\text{stage } 0}, \quad (6)$$

where $d_{ic(T)}$ is an indicator that equals 1 if the index spouse i shares information with their partner s , and zero otherwise, and δ is the weight put by the secondary spouse on the information provided by the index spouse.

To test the hypothesis that there is spillover of information from the index spouse to the secondary spouse, we will compare the secondary spouse stage 5 beliefs in the treatment and the control groups. More precisely we will test for equality in their average absolute perception gaps:

Hypothesis 2: $\overline{PG_{sc(T)}^{\text{stage } 5}} = \overline{PG_{sc(C)}^{\text{stage } 5}}.$

As we will show below, we find evidence that secondary spouses in treated couples have smaller perception gaps, on average, than those of their counterparts in control couples at stage 5, the first time we interview them: specifically, they are closer to the information that was given to the treated index spouses. This, in itself, is direct evidence of spillovers.

Note that no spillover may arise if the index spouse does not share any information ($d_{ic(T)} = 0$), and/or if the secondary spouse fully discounts the information provided by their spouse ($\delta = 0$). Standard economic theory suggests that the index spouse is more likely to share information when it is valuable and there are greater gains from sharing. In our context, the primary benefits of sharing arise when the information can be used to re-optimize decisions related to saving and labor force participation. However, information sharing may be less likely if there are associated costs. These costs may be psychological (e.g., the conversation could trigger conflict), temporal (e.g., it may take time to engage in the discussion), or cognitive (e.g., the complexity of the information may make it harder to explain or understand). Furthermore, strategic motiva-

tions may lead the index spouse to withhold information. For example, they may choose not to disclose negative news about their future benefits to preserve their bargaining power within the relationship. Some of these themes emerge in the qualitative data that we collected as open-ended questions in a follow-up survey; these are discussed in section 4.4.2. Finally, the secondary spouse may in turn discount some or all of the information provided by the index spouse for several reasons. They may question the accuracy of the information or distrust their spouse’s financial knowledge or interpretation.

To explore these dynamics further, we conduct heterogeneity analysis of the spillover effects below to better understand what drives information spillovers in some couples but not others. Our aim is to provide novel evidence on which types of couples experience greater belief revisions by the secondary spouse, indicating stronger information spillovers. Identifying the separate contributions of $d_{ic(T)}$ and δ falls outside the scope of this paper.

4.2 Treatment Effect on Index Spouses’ Social Security Benefits Expectations

We first investigate the causal effect of the information provision on belief-updating of the index spouse. Understanding this initial response is critical, as it serves as a necessary “first stage” to assess potential information spillovers to the secondary spouse. Without belief updates from the index spouse, it would be impossible to evaluate spillover effects. Our primary outcome of interest is the absolute perception gap for the age 67 benefits, which quantifies the deviation of the index spouse’s expectations from the information provided in the treatment, namely:

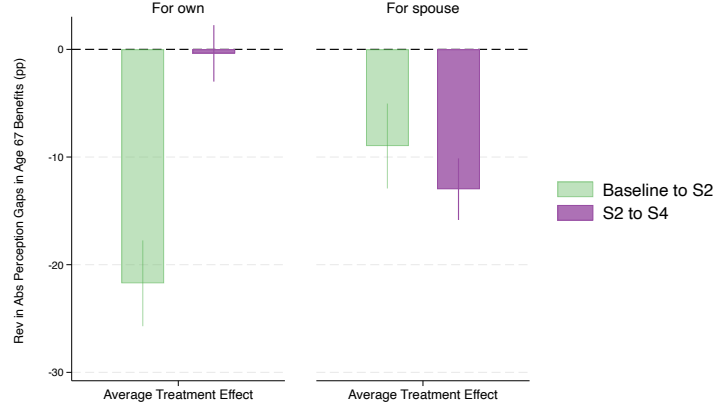
$$PG = \frac{|\text{Predicted benefits by the SS Calculator} - \text{Expected benefits}|}{\text{Predicted benefits by the Social Security Calculator}}. \quad (7)$$

We calculate this absolute perception gap for index spouses in both the treatment and control groups. In the control group, the predicted benefits are the ones that respondents would have received if they had been in the treatment group. A decrease in the absolute perception gap reflects a closer alignment of beliefs with the predicted benefits. We, therefore, focus on the revision in absolute perception gaps across the various survey stages, expecting a reduction in the absolute perception gap for treated respondents.

For index spouse i , we denote the revision of absolute perception gap between stage 0 and stage 2 by $\Delta PG_{i,2,0} = PG_{i,s2} - PG_{i,s0}$. We are interested in the absolute perception gap of a respondent for their own benefits ΔPG_i^{own} and those for their spouse ΔPG_i^{spouse} . We first examine the revision from stages 0 to 2 ($\Delta PG_{i,2,0}$) and then from stages 2 to 4 ($\Delta PG_{i,4,2}$); as shown in Figure B1, the former is the updating as a result of information about own Social Security benefits only, and the latter is the updating as a result of additional information about spouse’s Social Security benefits. We estimate the following specifications:

$$\Delta PG_{i,s,s-2}^h = \eta_1 + \beta_1 T_i + \theta_1 X_i + \epsilon_i, \quad s \in \{2, 4\}, h \in \{own, spouse\}, \quad (8)$$

FIGURE 4: Treatment effect on index Spouse’s absolute perception gaps



Note: The figure shows the average treatment effects from stage 0 to 2 (in green) and from stage 2 to 4 (in purple), estimated using equation (8), separately for index spouse benefits and secondary spouse benefits. The spikes on each bar show 95% confidence intervals. Appendix Figure A5 shows the average treatment effects from baseline (stage 0) to endline (stage 4). Appendix Table A6 presents the estimates.

where T_i is a dummy indicating if the index spouse i is in the treatment group, and X_i is a vector of index and secondary spouse’s characteristics, including education, age, ethnicity, gender, and financial literacy. We use heteroskedasticity-robust standard errors. In equation (8), η_1 measures the mean revisions in the control group and soaks up any effects attributable to the mere act of taking the survey or of reporting expectations multiple times. The parameter of interest is β_1 , which measures the treatment effect on the revision of the absolute perception gap.

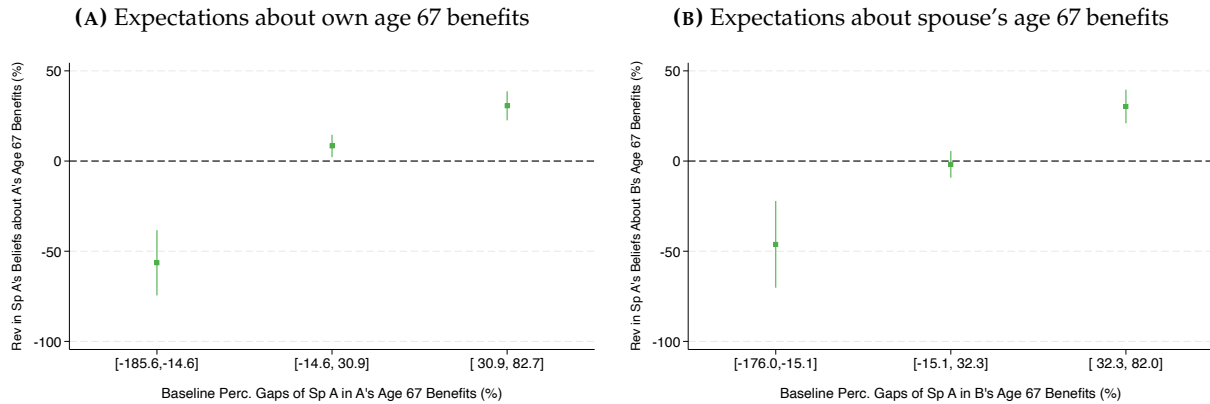
Figure 4 displays the treatment effects for expected benefits conditional on claiming at age 67. In the left panel, the green bars represent the treatment effect for own expected benefits from stages 0 to 2, while the purple bars represent stages 2 to 4. The right panel illustrates the analogous treatment effects on index spouse’s absolute perception gaps for the secondary spouse’s benefits. The figure reveals considerable revisions to expectations in the treatment group. The left panel shows that the stage 1 information treatment (on own benefits) had a large and precise effect on the revision in the absolute perception gap for own benefits, with a reduction of 21.7 percentage points (green bar), or 39% of the baseline mean. However, the stage 3 information treatment (on spouse’s benefits) shows no additional effect, with a treatment effect close to zero (purple bar). This is quite sensible: treated index spouses’ expectations about their own benefits become more aligned with the provided information on own benefits, with no further revision upon receiving information on spouse’s benefits.

The right panel shows that the index spouses revise their beliefs about their spouse’s benefits upon receiving information about both own and spouse’s benefits. Although the stage 1 information treatment was about own benefits, it led a reduction of 9 percentage points or 16% of the baseline mean in the absolute perception gap for spousal benefits (green bar). This suggests that treated index spouses perhaps gain a more comprehensive understanding of the Social Security

formula when initially learning about their own benefits. Moreover, the stage 3 information treatment specifically targeting spousal benefits leads to an additional reduction of approximately 13.1 percentage points or 23% of the baseline control mean. Consequently, the index spouses' expected benefits for their spouse become more closely aligned with the information provided. Overall, the revisions of absolute perception gap from stages 0 to 4 are of similar magnitude, about 22 pp, for own and spouse's expected benefits (Appendix Figure A5).

Figure 4 presents the average treatment effect (and Appendix Table A6 shows the corresponding estimates), but we anticipate heterogeneous treatment effects based on the baseline perception gap. Individuals with initial expectations lower (respectively, higher) than the provided information should update their expectations upward (respectively, downward). Figure 5 presents the treatment effect for the (Stage 4 - Stage 0) revision in index spouse's beliefs as a function of the baseline perception gap. As anticipated, index spouses who initially overestimated (negative perception gap) adjusted their beliefs downward (average treatment effect of -56 percentage points for own benefit expectations). Conversely, those who initially underestimated (positive perception gap) updated their beliefs upward, with a treatment effect of +30 percentage points for own Social Security benefit beliefs. There is a treatment effect of essentially zero for individuals whose baseline beliefs were already closely aligned with the provided information. Overall, the updating patterns are consistent with index spouses responding to the informational content in the signal that is given to them.

FIGURE 5: Heterogeneity in the treatment effect on index spouse's benefit expectations, by terciles of baseline perception gaps



Note: The figure shows the average treatment effects (from baseline to endline) interacted with terciles of baseline perception gaps, estimated using equation (8), separately for index spouse benefits and secondary spouse benefits. The spikes show 95% confidence intervals. Appendix Table A7 shows the average treatment effects interacted with terciles of the baseline perception gaps separately from stage 0 to 2 and from stage 2 to 4.

Appendix Figure A6 explores the heterogeneity in the treatment effect for index spouses for own benefit expectations based on observable characteristics (Appendix Table A8 shows the corresponding estimates). For this analysis, we estimate equation (8) for different subsamples with the relevant characteristics. A more negative estimate suggests a larger decline in the absolute

perception gap (that is, a larger treatment effect as the expectation becomes more aligned with the provided information). Looking across the various subsamples, we see that the estimate is less negative (i.e., a smaller treatment effect) if the index spouse is in a couple where at least one spouse is above the age of 55 or if the couple's total earnings are below median. This may be due to these groups being better informed (see Appendix Table A5).

We also see that the estimates are more negative (i.e., larger treatment effects) if the index spouse trusts the information or is the financial decision-maker in the household, or if they are negatively surprised (i.e., they learn that their benefits will be lower than what they had expected). While none of these patterns are significant at conventional levels, they are quite sensible.

4.3 Information Spillover to Secondary Spouses

Our main objective is to investigate whether information spills over to the secondary spouse. Specifically, we compare the *stage 5* beliefs of secondary spouses for treatment and control groups. This stage corresponds to the first elicitation of expectations from the secondary spouse, as highlighted in orange in the study design shown in Figure B1.

For a secondary spouse j in couple c with an absolute perception gap PG_{jcs}^h at stage s and demographic characteristics X_{jc} , we run the regression:

$$PG_{jcs}^h = \eta + \beta T_c + \theta X_{jc} + \epsilon_{jcs}, \quad s \in \{5, 7\}, \quad h \in \{own, spouse\}, \quad (9)$$

where T_c is a dummy indicating if a couple was treated (that is, the index spouse in the couple received information), and X_{jc} is a vector of individual/couple-level characteristics, including education, age, ethnicity, gender, financial literacy and numeracy skills of both spouses. We use heteroskedasticity-robust standard errors. The parameter of interest is β . In the absence of spillovers, this will be zero for $s = 5$, i.e., when only the index spouse has received information. However, in the presence of spillovers, we expect this to be negative: that is, we expect the absolute gaps of secondary spouses in treated couples to be smaller than those of secondary spouses in control couples.

Table 5 shows the estimates of this regression for $s = 5$. The dependent variable in the first two columns is the absolute perception gap in percentages for secondary spouse's beliefs about index spouse's and own benefits, respectively. We see that baseline (i.e., Stage 5) perception gaps of secondary spouses in treated couples are smaller by about 10 to 13 percentage points in both cases. Given that the baseline absolute perception gaps of control secondary spouses is about 53 to 55 percent, this is a sizable decline.

The average decline of 10 to 13 percentage points is especially notable, given that the mean revision of the absolute perception gap for index spouses is about 22 percentage points (Appendix Figure A5). The estimate suggests that nearly half of the treatment is transmitted to the secondary spouse, indicative of significant spillover of information within couples.

TABLE 5: Information spillovers to the secondary spouse

| | Percents | | Dollars | |
|-----------------------|----------------------|--------------------|----------------------|---------------------|
| | (1) Index Sp Ben. | (2) Sec Sp Ben. | (3) Index Sp Ben. | (4) Sec Sp Ben. |
| Treated Couple | -12.8*** (2.4) | -12.6*** (2.7) | -318.4*** (63.4) | -299.2*** (69.1) |
| TE as % of CG Abs PG | -24.0 | -22.9 | -26.5 | -25.2 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ |
| Control Dep. Var Mean | 53.3 | 55.1 | 1199.3 | 1189.5 |
| Adj. R-Squared | 0.0 | 0.1 | 0.1 | 0.1 |
| Observations | 2,029 | 2,006 | 2,029 | 2,006 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in the first two columns is the absolute perception gap of the secondary spouse in index spouse's or their own age 67 benefits as share of the predicted benefits by the SS Quick Calculator. The dependent variable in the last two columns is the absolute perception gap of the secondary spouse in index spouse's or their own age 67 benefits in dollars. Demographics include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy for both the index and the secondary spouses.

This estimate is likely biased downwards since survey beliefs are reported with error. In a simulation exercise, we find that the spillover would be much higher (with a median equivalent to 70% and a 90th percentile equivalent to 82% of the index spouse treatment effect, out of 1000 bootstrap repetitions) if there was full information-sharing and beliefs were reported with measurement error, suggesting that the incomplete sharing of information between spouses cannot be fully explained by measurement error alone.¹⁸

These results point to the presence of information spillovers, but also highlight substantial frictions in the flow of information within couples. To benchmark the magnitude of the observed spillovers, we compare it to the belief updating that occurs when the secondary spouse receives the information directly. This occurs in treatment T1, where the secondary spouse is presented with the same information at stage 6 that the index spouse received earlier (see section 2.3). We re-estimate the regression in Equation (9) using stage $s = 7$ beliefs (post-information) and report the results in Table 6. We find that directly informing the secondary spouse reduces the absolute perception gap by approximately 28 percentage points compared to the control group, more than twice the size of the spillover effect estimated in Table 5. This treatment effect serves as an upper bound for what information transmission could achieve in the absence of frictions.

The fact that the spillovers fall far short of this upper bound provides strong evidence that there are substantive barriers to effective information flow between spouses. These frictions may

¹⁸To assess the potential role of measurement error, we assume full information sharing between spouses but that expectations are reported with error. We first assume the secondary spouses' stage 5 expectations are equal to the index spouses' stage 4 expectations, but that they are reported with classical measurement error, modeled as normally distributed with a standard deviation of 0.35, which corresponds to the standard deviation of the test-retest classical measurement error of the control index spouses as discussed in Section 3. We re-estimate equation (9) 1,000 times with different draws for the measurement error from the same distribution and obtain larger spillovers 905 times for index spouse benefits and 987 times for secondary spouse benefits. The histograms showing the distribution of spillovers as a share of the index spouse treatment effect are presented in Appendix Figure A7.

TABLE 6: Secondary spouse average treatment effect - T1

| | Percents | | Dollars | |
|-------------------------|----------------------|--------------------|----------------------|---------------------|
| | (1) Index Sp Ben. | (2) Sec Sp Ben. | (3) Index Sp Ben. | (4) Sec Sp Ben. |
| Treated Couple | -27.4*** (3.0) | -28.3*** (3.2) | -663.5*** (79.3) | -640.1*** (78.2) |
| TE as % of CG S7 Abs PG | -48.6 | -48.7 | -51.9 | -52.2 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ |
| Control Dep. Var Mean | 56.3 | 58.0 | 1277.3 | 1225.6 |
| Adj. R-Squared | 0.1 | 0.1 | 0.1 | 0.1 |
| Observations | 961 | 954 | 961 | 954 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in the first two columns is the stage 7 absolute perception gap of the secondary spouse in index spouse's or their own age 67 benefits as share of the predicted benefits by the SS Quick Calculator. The dependent variable in the last two columns is the stage 7 absolute perception gap of the secondary spouse in index spouse's or their own age 67 benefits in dollars. Demographics include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy for both the index and the secondary spouses.

arise from limited communication, strategic considerations, or cognitive barriers, which we discuss below. Importantly, this pattern allows us to reject the idea that most couples fully share information but use it differently because, for example, they have different beliefs about the underlying data generation process.¹⁹ If that were the case, the direct and indirect treatment effects would be of similar magnitude. Instead, the large gap between them suggests that information fails to circulate effectively within households.

4.4 Heterogeneity in Spillovers

We next explore the heterogeneity in spillovers within couples to better understand which types of couples experience larger information spillovers.

4.4.1 Heterogeneity Analysis

Figure 6 presents the heterogeneity analysis for the index spouse's expected benefits, with corresponding regressions in Appendix Tables A10-A14, panel A. As before, a larger negative treatment effect on the secondary spouse baseline absolute perception gap indicates greater spillovers. Appendix Figure A8, focusing on spillovers about the secondary spouse's benefits, shows overall similar patterns.

Value of Information: Standard economic theory suggests that valuable information is more likely to be shared. We explore whether information spillovers are greater in couples for whom the information holds higher value. We begin by considering treatment-related factors. Panel A of Figure 6 investigates heterogeneity by whether the index spouse finds the provided information

¹⁹For instance, one spouse might anticipate a collapse of the Social Security system while the other does not.

credible.²⁰ About 60% of the index spouses find the information credible. Such individuals are precisely the ones who are leading to spillovers: in such couples, the absolute perception gap is lower by more than 18 percentage points versus by less than 4 percentage points for their counterpart couples. Additionally, spillovers are five times larger among couples where the index spouse's initial beliefs were less accurate (i.e., further from the predicted benefits presented), and for whom we expect the information to be more useful. The pattern and magnitude are the same for spillovers regarding the secondary spouses' benefits (see Figure A8).

Next, we explore the characteristics of couples for whom the information is likely more valuable — particularly those less familiar with the topic and with more time to adjust their savings or labor force participation, such as couples farther from retirement. Indeed, the figure shows that spillovers are primarily driven by couples under the age of 55, where the perception gap is lower by 16 percentage points. Additionally, the information on the index spouse's benefits is likely to be more valuable when both partners intend to claim on the index spouse's record. Indeed, spillovers are around 50% larger when the secondary spouse plans to claim on the index spouse's record; however, this difference is not statistically significant at conventional levels (p-value of 0.13).

Overall, this aligns with the theoretical argument that information with higher perceived value is more likely to spill over from one spouse to the other.

Strategic Motivations: We next explore whether strategic motivations play a role in how information is shared. This relates to a literature, primarily from low-income countries, on the strategic concealment of information from spouses (e.g., Chen, 2013; de Laat, 2014; Ashraf, 2009). In situations of asymmetric information, there is typically a trade-off between maximizing individual gains and maximizing overall efficiency. In our context, an individual might seek to maximize personal utility by concealing information that weakens their bargaining position within the couple (e.g., discovering their own Social Security benefits are lower than expected, or their partner's benefits are higher).²¹ However, withholding such information can be inefficient. For instance, sharing negative income shocks (e.g., that their own or their spouse's benefits are much lower than they had expected) may be beneficial, as failing to re-optimize the couple's choices in light of this information could have significant costs (especially when utility over consumption is concave). In contrast, the cost of not re-optimizing may be lower if the new information is positive (e.g., that benefits are higher than expected).²²

The first two cuts in Panel B of Figure 6 investigate if the nature of spillovers differs by whether

²⁰The question was: "On a scale of 1 to 5 (where 1 is "Not At All" and 5 is "Fully"), how much do you trust this information?" An index spouse who selects 4 or 5 on this scale is coded as having found the information credible.

²¹Conversely, they might be more likely to share information that strengthens their position (e.g., that their own benefits exceed expectations or their partner's benefits fall short).

²²Individuals who experience a negative shock are more likely to have under-saved relative to the optimal saving-consumption path, while those who experience a positive shock are more likely to have over-saved. A priori, it is unclear whether under-saving or over-saving leads to greater welfare losses. However, for most agents with concave utility, bequest motives, borrowing constraints, and patience, under-saving is expected to be more detrimental. It can result in sharp consumption drops, unfulfilled bequest intentions, and limited flexibility to adjust consumption due to borrowing constraints.

the index spouse received positive news (i.e., learning their predicted Social Security benefits are higher than their expectations) or negative news about their benefits. We find significantly larger spillovers when the news about the index spouse's benefits is negative. Similarly, spillovers are also greater when the index spouse learns negative information about their partner's Social Security benefits. The average reduction in the absolute perception gap is about 19-22 percentage points for negative news as opposed to 7-8 percentage points for positive news. These patterns are consistent with efficient sharing of information within the household, where individuals are more likely to share when the cost of not re-optimizing decisions is likely higher. It is also possible that negative news, which tend to be more salient, are more likely to be shared; some papers in the literature on asymmetric belief-updating tend to find more responsiveness to negative news (Benjamin, 2019).

To further test for the strategic motivations, we compare spillovers between more equal and less equal couples—specifically, whether both spouses are working or if there is a large difference in earnings. The rationale is that income or labor market differences between partners may proxy for differences in bargaining power within the household (Lundberg and Pollak, 1996), potentially leading to a greater motivation to hide in less-equal couples. However, we find no significant difference in information spillovers, suggesting that strategic motivations play a minimal role in sharing information about Social Security benefits.

Communication Frictions: Next, we examine potential costs that may hinder communication. These costs can arise from concerns of sparking conflict, stress from discussing financial topics, or diverging interest, as pointed out in our open-ended questions (see discussion in section 4.4.2). Panel C presents various analyses based on proxies for the absence of communication frictions within couples, such as agreement on the quality of their marital relationship, alignment in beliefs about total household earnings, and agreement on whether the index spouse handles financial decisions. Agreement on these factors likely indicate regular and effective communication, suggesting lower communications costs in such couples.

A couple is coded as being in less agreement if the difference between the marital satisfaction index for each spouse is above the sample median or if the difference in the two spouses' perception of the total household earnings is above the median. In each of the heterogeneity analyses in Panel C, spillovers are (weakly) larger when the couple is in agreement. For example, the average reduction in the absolute perception gap of the secondary spouse is 19.0 percentage points in couples who are in more agreement regarding marital satisfaction versus 7.4 percentage points for those who are in less agreement. It is worth noting that the level of marital satisfaction is not a predictor of larger spillovers (see Panel E). This suggests that it is couples agreeing on the quality of the relationship (and not the quality level per se) that matters for information spillovers.²³

²³We also examine the time gap between when the index and secondary spouse complete the survey. A longer gap may provide the index spouse with more time to share information, but it could also lead to weaker recall. While we do not have exogenous variation in this time gap, splitting the sample into below- and above-median gaps reveals that spillovers are in fact larger when the time gap is shorter. It is hard to interpret this finding since couples with smaller gaps are not a random subset (for example, they tend to have below-median marital length).

Cognitive frictions: We analyze potential cognitive costs, as the complexity of the Social Security system may pose challenges. Effective information transmission depends on the index spouse’s ability to understand and communicate the information, and the secondary spouse’s ability to process it. Couples with higher education or financial literacy are likely better equipped to handle this process.

Panel D of Figure 6 shows heterogeneity based on couples’ education, financial literacy, and the index spouse’s comprehension of the information, as measured by follow-up questions. Overall, there is little evidence that cognitive frictions significantly influence information spillovers, as no clear patterns in treatment effects are observed across these groups.

Couple-based factors: Panel E concludes the analysis by investigating whether certain characteristics of the index spouse or couple influence information spillovers. Prior work has found evidence of larger spillovers in couples when the information is provided to the male (Conlon et al., 2022; Fehr et al., 2023). Spillovers are somewhat larger when the index spouse is male—reducing the perception gap by 16 percentage points compared to 10 percentage points when the index spouse is female—but this difference is not statistically significant ($p = 0.24$). The result is similar when looking at the secondary spouse Social Security benefits (Appendix Table A14, Panel D).

We also find that spillovers are larger among couples with children under the age of 18 and among those with above-average net worth, suggesting that family structure and financial resources may influence how information is shared and internalized.

4.4.2 Alternative Interpretations

We now explore alternative interpretations for the results of our heterogeneity analysis.

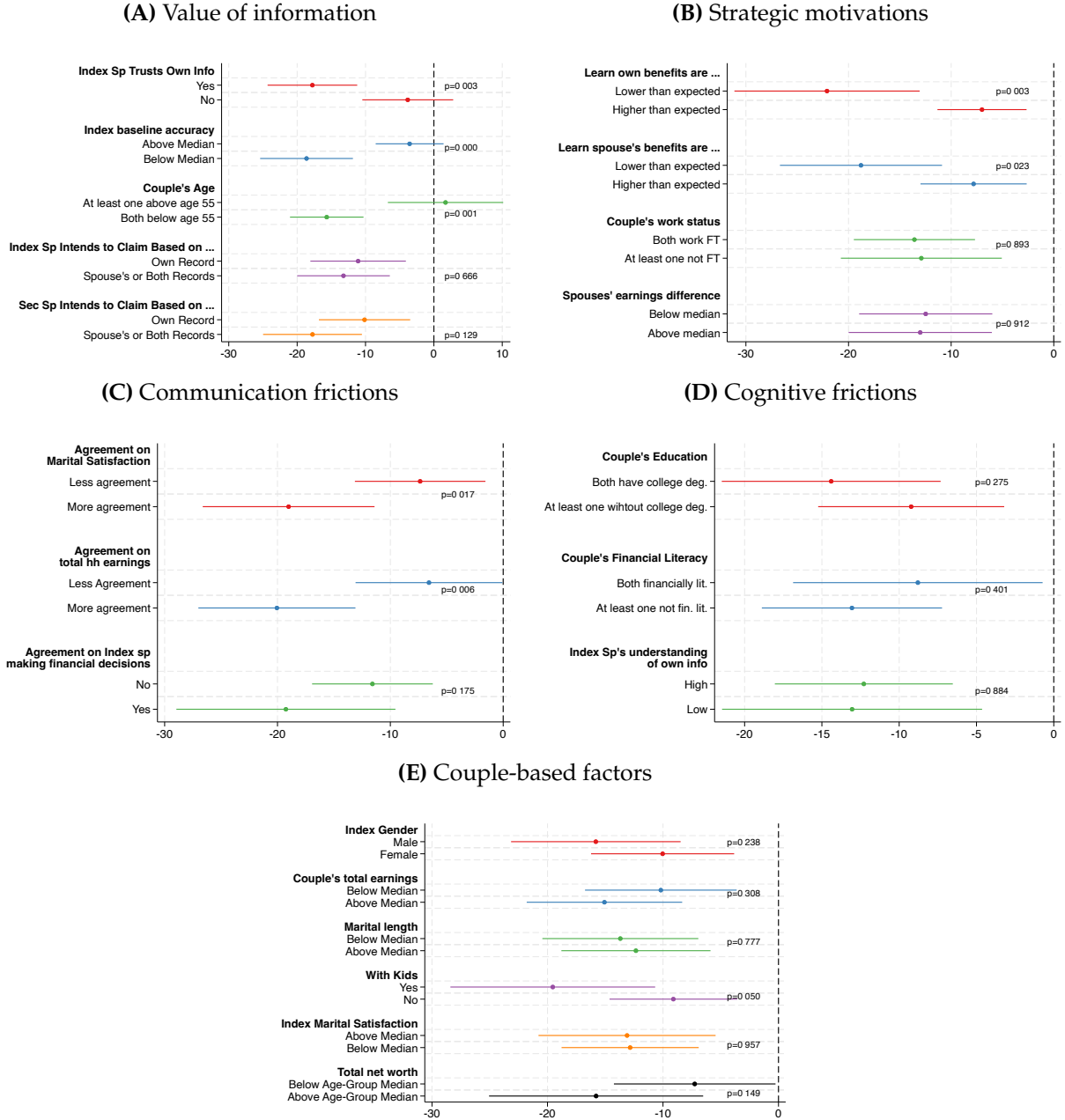
First, we investigate whether the observed smaller impact on perception gap for some secondary spouses is due to the index spouse revising their expectations less following the information treatment, rather than actual lower spillovers of information between spouses. To address this, Appendix Figure A9 (and corresponding Appendix Tables A10-A14, panel B) presents the heterogeneity analysis of spillovers in the index spouse’s benefits while controlling for the index spouse’s baseline perception gap. The rationale for adding this control is that index spouses with larger perception gaps tend to update their expectations substantially more.²⁴

The main effects are largely unchanged when controlling for the index spouse baseline perception gap (see Appendix Table A9). Moreover, with this new specification, the patterns remain identical to those described earlier, suggesting that the heterogeneity arises from processes occurring *after* the index spouse revises their beliefs—either they have not communicated the information, or the secondary spouse has not updated their expectations despite the communication.

Second, we examine whether the larger spillovers are mechanical, specifically whether they

²⁴We prefer not controlling for the index spouse’s *revision* in perception gap as it is not pre-determined with respect to treatment (and is in fact endogenous), but note that the main results and heterogeneity analysis are similar when controlling for it.

FIGURE 6: Heterogeneity in information spillovers to the secondary spouse



Note: Figures show the spillovers in index spouse benefits estimated using equation (9) separately for different groups. Spikes show 95% confidence intervals and p-values are for testing difference in spillovers across mutually exclusive groups. Appendix Figure A8 shows the spillovers in secondary spouse benefits for the same groups.

are driven by groups with initially larger perception gaps, giving them more margin to adjust their expectations. This could change the interpretation of our findings. For instance, do younger couples have larger spillovers because the information is inherently more valuable to them, or because their initial perception gap is wider? To investigate this, we replicate our heterogeneity

analysis using a within-group normalized perception gap as the dependent variable. Specifically, for each subgroup (e.g., younger vs. older couples), we calculate a z-score for the perception gap, with a mean of zero and a standard deviation of one. Appendix Tables A10-A14 panels C and F show the heterogeneity analysis in spillovers with this adjusted dependent variable. Our conclusions remain largely unchanged. For example, we still find significantly larger spillovers among couples where both spouses are under the age of 55 compared to older couples, or among couples whose index spouse's baseline perception gap in own benefits was below the median. Estimates for the subgroup receiving negative news are less precise but remain qualitatively consistent.

4.4.3 Dominant Factors in Information Spillovers and Multiple Hypothesis Testing

As we explore heterogeneity along many different characteristics, we face issues of multiple hypotheses testing. To mitigate this concern, we create summary indices by combining variables for each of the factors shown in panels (A)-(D) of Figure 6 so that we effectively test only 4 hypotheses. Panel E variables are directly controlled for in this analysis, since it is not clear how to create a weighted index of couple-based factors.

For ease of interpretation, the indices are each standardized with mean 0 and standard deviation 1, and are constructed as follows:

- **Lack of Cognitive Barriers:** This index is based on the following components: (i) both spouses have a college degree, (ii) the index spouse is financially literate, and (iii) the index spouse shows high comprehension in the information understanding checks (after receiving information about their own or their spouse's benefits). We sum these indicators and standardize the total.
- **Lack of Communication Frictions:** This index includes: (i) below-median disagreement in marital satisfaction, (ii) below-median disagreement in reported household earnings, and (iii) agreement that the index spouse is responsible for household financial decisions. These components are summed and the total is standardized.
- **High Value of Information:** This index adds the variable from panel (A) of Figure 6 with one point for each of the following: (i) at least one spouse is under age 55, (ii) the index spouse's expected age 67 benefits are less accurate (i.e., below-median accuracy), (iii) the index spouse intends to claim based on the spouse's record, and (iv) the secondary spouse intends to claim based on the spouse's record. The sum is then standardized.²⁵
- **Negative News:** This index is based on whether the benefit information received is less favorable than initially expected. For index spouse spillovers, this reflects whether their projected benefits are lower than expected. For secondary spouse spillovers, it reflects whether

²⁵When analyzing spillovers in secondary spouse benefit expectations, we use the secondary spouse's benefit accuracy instead of the index spouse's.

the spouse's projected benefits are lower than initially expected. The indicator is then standardized.

- **Equal Couples:** This index includes one point each if (i) both spouses work full-time, and (ii) the earnings differential between them is below the median. The total is then standardized.

We bring all this together in Table 7 where the absolute perception gap for the secondary spouse is regressed onto an indicator for whether the index spouse was treated, and its interaction with the various indices (the indices are also included by themselves). We show the results in columns 1 and 2 for the index spouse's benefits and 3 and 4 for the secondary spouse's benefits. Columns 2 and 4 control in addition for the couple-based factors as in panel (E) of Figure 6.

The table reveals that the primary drivers of spillovers are the absence of communication frictions, a high value of information and receiving negative news. A one standard deviation increase in the lack of communication frictions index leads to spillovers related to the index spouse's benefits that are 55% larger and spillovers related to the secondary spouse's benefits that are doubled, compared to couples with a zero value for the index. Similar effect sizes are observed for the High Value of Information and Negative News indices in the case of index spouse expectations, with somewhat smaller effects for the secondary spouse. Given that sharing negative news likely yields greater informational benefits than sharing positive news, our results suggest that the value of information plays a crucial role in driving spillovers.

Overall, our results are consistent with cooperative models of decision-making with information frictions. Notably, despite focusing on a cognitively complex decision, our findings suggest that these frictions are not cognitive in nature but instead stem from barriers to communication.

4.4.4 Qualitative Evidence from the open-ended questions

To complement the findings from our analysis, we included several open-ended questions in a follow-up survey, focusing on within-couple communication. This approach aligns with recent work in economics that uses open-ended questions to explore what is top of mind for individuals (for example, see Stantcheva et al., 2024). 153 couples were successfully re-interviewed between May 29 and July 31, 2024, about 27 months after the first survey. These respondents look largely similar to our original sample. Specifically, respondents were asked: *"Tell us a little bit about how often you discuss financial issues with your spouse. Specifically, what kinds of issues do you discuss, who has more information typically, who ends up taking advice from the other spouse, etc."*

These open-ended responses are consistent with our heterogeneity analysis of spillover effects. Couples frequently emphasize that financial discussions revolve around managing savings and planning for retirement, suggesting that information on Social Security benefits could be particularly valuable. Additionally, many responses indicate joint decision-making in this area, highlighting the importance of sharing information between partners. However, the responses also point out key barriers to communication. Stress and fear of conflict are prominent reasons for

TABLE 7: Factors in spillovers to the secondary spouse

| | Index Sp Benefits | | Sec Sp Benefits | |
|---|-------------------|------------------|------------------|------------------|
| | (1) | (2) | (3) | (4) |
| Treated | -11.6*** (2.6) | -13.9** (6.4) | -9.8*** (2.9) | -10.2 (6.8) |
| Treated \times Lack of cogn barriers | 0.6 (2.1) | 1.1 (2.3) | 3.5 (2.4) | 3.2 (2.8) |
| Treated \times Lack of comm frictions | -6.4*** (2.3) | -6.4*** (2.4) | -7.6*** (2.5) | -8.2*** (2.6) |
| Treated \times High value of info | -5.5** (2.3) | -5.4** (2.4) | -5.4** (2.7) | -5.3* (2.7) |
| Treated \times Negative news | -5.2** (2.5) | -5.6** (2.5) | -6.6** (2.8) | -7.5*** (2.9) |
| Treated \times Equal couples | 2.6 (2.3) | 2.0 (2.3) | -0.5 (2.5) | -1.0 (2.5) |
| Couple-based Factors | | ✓ | | ✓ |
| Control Dep. Var Mean | 53.3 | 53.3 | 55.1 | 55.1 |
| Adj. R-Squared | 0.1 | 0.1 | 0.1 | 0.1 |
| Observations | 2,029 | 2,029 | 2,006 | 2,006 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in all columns is the absolute baseline perception gap as a share of predicted age 67 benefits by the Social Security quick calculator. The specifications also include the uninteracted indices. Robust standard errors are in parentheses. Appendix Table A15 shows the estimates when we also control for index spouse's baseline perception gaps in own or spouse's benefits. Demographics include the ethnicity and perception of own financial savviness of both the index and the secondary spouses.

avoiding financial discussions. Furthermore, the fact that many couples discuss finances on a weekly or monthly basis suggests the existence of a fixed cost in initiating such conversations. We summarize the findings in greater detail, together with word clouds in Appendix C.

5 Alignment of Intended Claiming Ages

We have shown that there are substantial spillovers of the provided information to the secondary spouse, and that there is meaningful heterogeneity in the spillovers. We now examine how information transmission (or lack thereof) within couples affects intra-household decision-making. To investigate this, we focus on the alignment in expected claiming ages for a given spouse, as discussed in Section 3.3. While we do not know the optimal claiming ages for couples, it must be the case that couples are not optimizing their current saving and labor supply decisions if one spouse expects the other spouse to retire and/or claim much later or earlier than what is intended by the spouse themselves.

Before turning to alignment, we first examine whether the information provision leads index spouses to update their expected claiming ages. Specifically, we estimate the same specification as in Equation (8), but with revision in index spouses' expectations for their own and for their spouses' claiming ages as the dependent variable, respectively. Appendix Table A16 shows that index spouses revise their expected claiming age upward by 0.35 years (around 4 months) and

their expectations about their spouses' claiming age by 0.45 years (around 5.4 months) after receiving the stage 1 and stage 3 information treatments.

Next, we consider alignment in spouses' expectations. Our measure of alignment is the absolute distance between spouses' reported expected claiming ages for the *same* spouse. For example, the absolute value of $|\text{index spouse's own claiming age expectation at stage 0} - \text{secondary spouse's claiming age expectation for index spouse at stage 5}|$ is the absolute distance in index and secondary spouses' beliefs about index spouse's claiming age at the baseline of their respective surveys (i.e., before they directly receive any information). We denote this as $D_{c,0,5}$. Likewise, $D_{c,4,5}$ is the absolute distance between the index spouse's updated expectation at stage 4 (after receiving information) and the secondary spouse's baseline expectation at stage 5. The change in alignment due to the information treatment to the index spouse is captured by $D_{c,4,5} - D_{c,0,5}$. A decrease in this measure indicates improved alignment as a result of information provision and/or spillovers.

We run the same specification as in equation (9) except that this misalignment measure is the dependent variable and that we cluster standard errors at the couple level.²⁶ Only T1 and Control couples are used for this analysis.²⁷ Estimates are presented in the first column of Table 8. We observe an average decrease of 1.1 months in misalignment when the index spouse is provided with information. When we restrict our sample to couples who had misalignment in their baseline expectations, in column 2, we find an economically and statistically decrease in misalignment of 3.4 months or by 8.7% of the control group baseline average. This is a meaningful effect and suggests that some efficiency gains are possible (that is, misalignment can decrease) when only one spouse is targeted with information.

An important question is whether additional gains—specifically, further reductions in misalignment—can be achieved when both spouses receive the information. To explore this, we leverage the T1 treatment, in which the secondary spouse was also provided with the same information as the index spouse at stage 6. The dependent variable of interest is $D_{c,4,7} - D_{c,0,5}$, i.e., the change in misalignment when both spouses have received and processed the information (stage 4 for the index spouse and stage 7 for the secondary spouse, as shown in Figure B1). Columns 3 and 4 of Table 8 show a reduction in the absolute distance, significant for couples with misaligned beliefs at the baseline, implying that providing information to both spouses closes the baseline misalignment gap by around 5.2 months or by around 13.4% of the control group baseline average. Even though the estimates in columns 2 and 4 are not statistically different (p-value=0.17), this suggests that informing both spouses may lead to weakly greater alignment in expected claiming age compared to informing only one spouse, even in the presence of information spillovers.

²⁶Here, the dependent variable is a vector of the absolute distance between spouses' reported expected claiming ages for both the index and the secondary spouse. This implies that, for couples with non-missing information, there will be 2 observations per couple.

²⁷We exclude secondary spouses in T2 since they received a different information treatment. This restriction allows us to compare the effect of providing the same information to one spouse versus both spouses in the same sample.

TABLE 8: Alignment in beliefs about Social Security benefit claiming age

| | Only Index Sp Receives Info | | Both Sp Receive Info | |
|--------------------------|-----------------------------|--|-------------------------|--|
| | (1) | (2) | (3) | (4) |
| | $D_{c,4,5} - D_{c,0,5}$ | $D_{c,4,5} - D_{c,0,5} D_{c,0,5} \neq 0$ | $D_{c,4,7} - D_{c,0,5}$ | $D_{c,4,7} - D_{c,0,5} D_{c,0,5} \neq 0$ |
| Treatment | -1.1 (1.3) | -3.4** (1.6) | -1.8 (1.5) | -5.2*** (2.0) |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ |
| Control Gr Baseline Mean | 24.1 | 38.9 | 24.1 | 38.9 |
| Adj. R-squared | 0.01 | 0.01 | 0.00 | 0.01 |
| Observations | 2,683 | 1,692 | 2,683 | 1,692 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in the first column is the difference in the absolute distance between index spouse's stage 4 and secondary spouse's stage 5 claiming age expectations and in the second column it's the same difference as in column 1, conditional on the absolute distance between index spouse's stage 4 and secondary spouse's stage 5 claiming age expectations being different than 0. In the third column, the dependent variable is the difference in the absolute distance between index spouse's stage 4 and secondary spouse's stage 7 claiming age expectations and the absolute distance between index spouse's stage 0 and secondary spouse's stage 5 claiming age expectations and in the last column it is the same difference as in column 3, but conditional on the absolute distance between index spouse's stage 0 and secondary spouse's stage 5 claiming age expectations being different than 0. Standard errors are clustered at the couple level. Demographics include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy for both the index and the secondary spouses.

6 Conclusion

Contrary to traditional economic models assuming perfect information flow within couples, we find significant misalignment in expectations about Social Security benefits. In our sample of US middle-aged couples, the correlation of Social Security expectations is 0.70. In addition, individuals within a couple tend to be misinformed in the same direction, with little evidence of specialization in information acquisition. This suggests that informational errors may be amplified, rather than corrected, within households. This pattern extends beyond Social Security. We find similar correlations, ranging from 0.46 to 0.64, for expectations concerning unemployment, inflation, stock market prices, household spending growth, and earnings growth. These descriptive findings are valuable in their own right, as we still know relatively little about how expectations align within households, which are a fundamental unit of economic decision-making.

We also provide causal evidence on information spillovers within couples by leveraging a randomized information intervention combined with a sequential survey design within couples. Consistent with some information sharing, having a treated spouse leads to a secondary spouse having expectations more in line with the information provided to the index spouse. The spillover effect is about half the treatment effect we estimate for the index spouse. The size of this spillover depends primarily on the gains from sharing and communication frictions within couples, with little evidence of strategic motives to hide (or share). Importantly, we also find that the information treatment enhances conditions for better intra-household decision-making.

Our paper is one of the first to document drivers of information spillovers within couples. While the economics literature has extensively studied how couples allocate time and money, less attention has been given to how they manage information, an equally valuable resource. Under-

standing the factors that drive the diffusion of information within couples is a first-order question. In addition, the limited literature on this topic has primarily focused on gender dynamics, leaving broader patterns of information transmission largely unexplored. Our findings reveal rich and intuitive patterns of heterogeneity, but further research is needed to understand why these patterns arise and how generalizable they are. For instance, while we find a limited role for cognitive frictions in our setting, it remains unclear whether this finding extends to other contexts. Additionally, our design cannot disentangle whether imperfect spillovers stem from the index spouse not communicating information or the secondary spouse discounting it. That would understandably be hard to disentangle in field settings; a stylized (laboratory) setup may be better suited to address that question and advance our understanding.

From a modeling perspective, our findings are consistent with cooperative models of household decision-making in the presence of information frictions. To advance the economics of the family, formal models should move beyond the assumption of complete information and incorporate heterogeneity in expectations, just as collective models allow for heterogeneity in preferences. One natural extension would be to explicitly model communication between spouses as a frictional process.

Our findings also carry important practical implications for the design of information interventions. Intra-household frictions can limit the effectiveness of policies that target only one member of the household. Interventions that not only provide information but also promote communication between spouses may enhance information transmission and improve joint decision-making. In the context of Social Security, for example, synchronizing benefit statements or reminders so that both spouses receive them at the same time—even well before retirement—could possibly facilitate communication and planning. More broadly, our results underscore the need to incorporate household dynamics into the design and evaluation of financial information policies.

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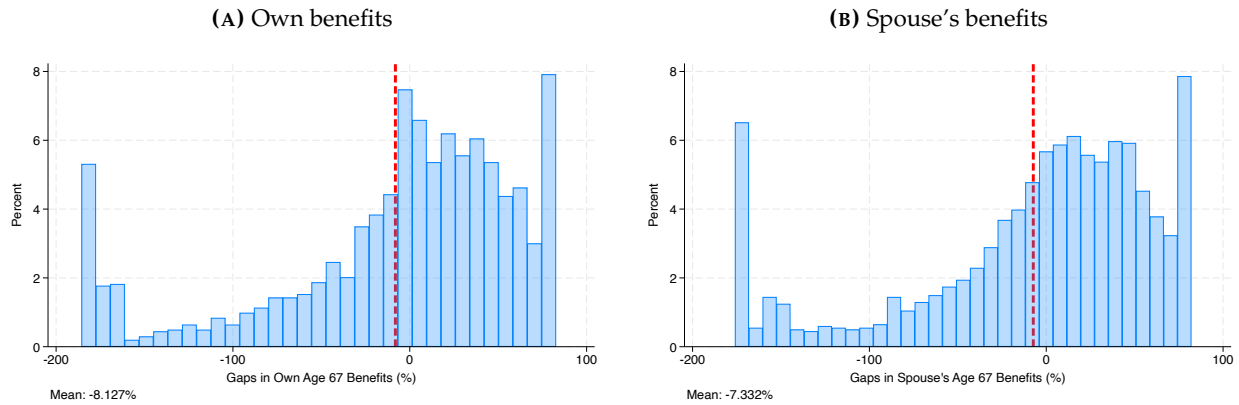
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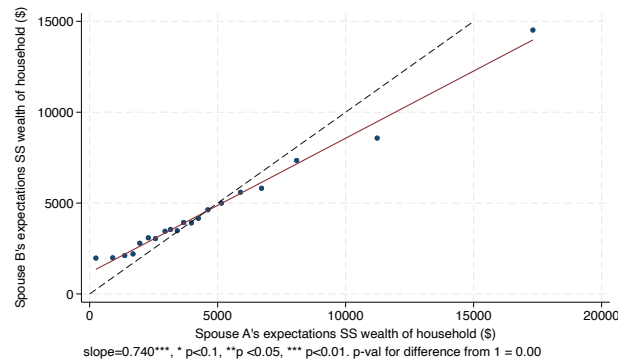
Appendix A Figures and Tables

FIGURE A1: Perception gaps in own and spouse's age 67 benefits



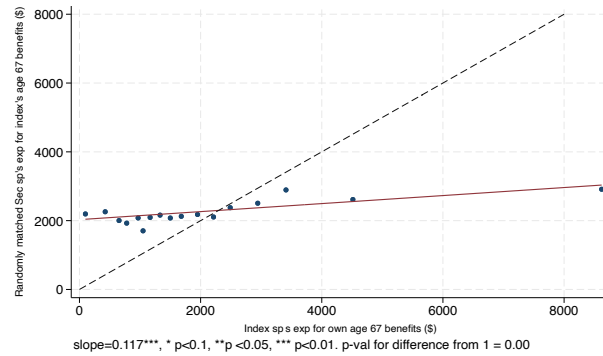
Note: The figures show the histograms for respondents' perception gaps in their own (left panel) and in their spouses' (right panel) age 67 benefits. Gaps are defined as $\frac{\text{Predicted benefits by the SSA calculator} - \text{Expected benefits}}{\text{Predicted benefits by the SSA calculator}}$. Gaps are winsorized at the top and bottom 5%. The red, vertical, dashed lines represent the mean of respondents' own (left panel) and spouses' (right panel) perception gaps.

FIGURE A2: Within-couple expectations for Social Security wealth



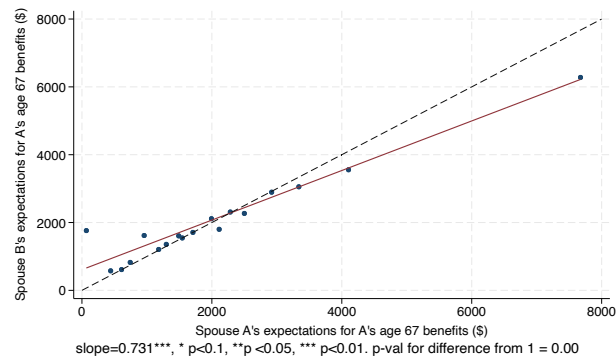
Note: The correlation between spouses' expectations for household's SS wealth is 0.74 (same as the regression slope shown in the figure) and this correlation falls to 0.68, when we exclude couples with SS wealth expectations above \$8,000. The correlation coefficient between the two variables used in the binscatter is 0.74, which is statistically different than 1 (p-value=0.00).

FIGURE A3: Correlation between spouses' benefit expectations - Randomly matching spouses



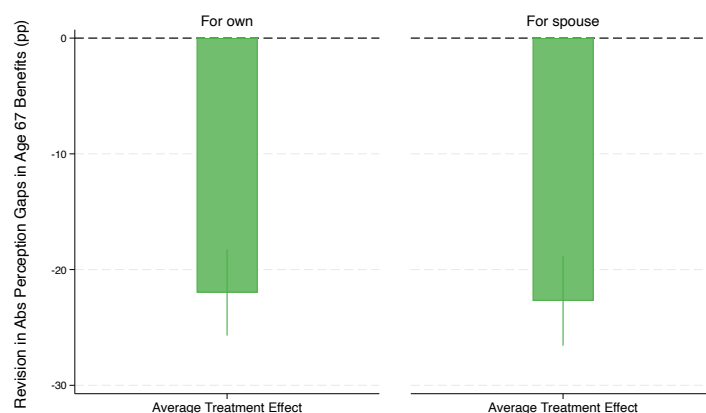
Note: The correlation coefficient between spouses' expectations for a given spouses' age 67 benefits falls to 0.12, when we randomly match secondary spouses to primary spouses based on the index spouses' characteristics (gender, age, education and employment status). This correlation is statistically different than 0.70 (p-value=0.00), which is the observed correlation coefficient between spouses' expectations in the data.

FIGURE A4: Correlation between spouses' benefit expectations - Couples with accurate earnings perceptions



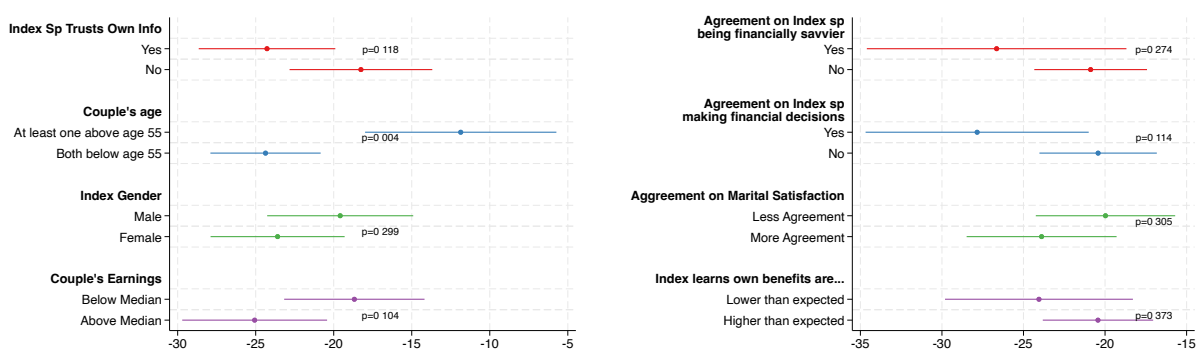
Note: The correlation coefficient between spouses' expectations for a given spouses' age 67 benefits is 0.75, when we restrict our sample to couples with accurate earnings perceptions. This correlation is statistically different than 1 (p-value=0.00).

FIGURE A5: Treatment Effect on index spouse's absolute perception gaps



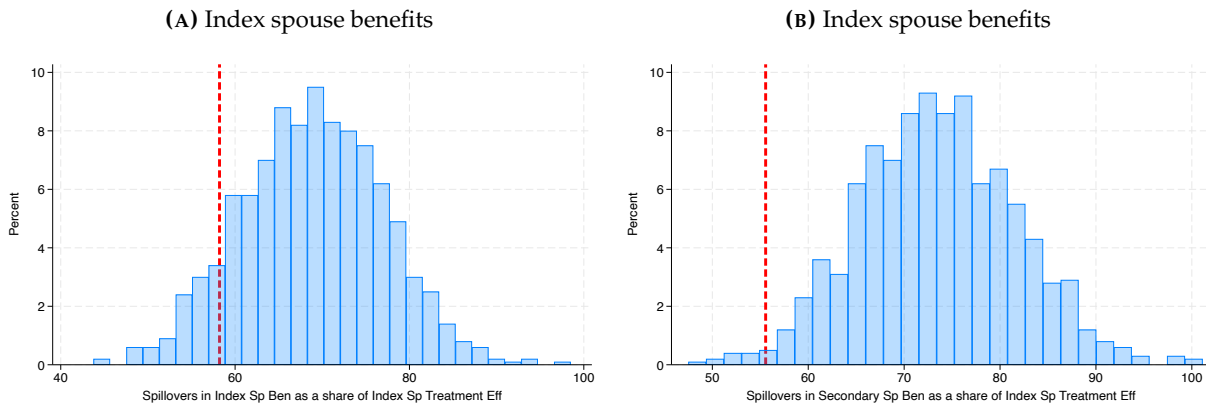
Note: The figure shows the average treatment effects from baseline to the endline, estimated using equation (8), separately for index spouse benefits and secondary spouse benefits.

FIGURE A6: Heterogeneity in the treatment effects on index spouse's absolute perception gaps about own age 67 benefits



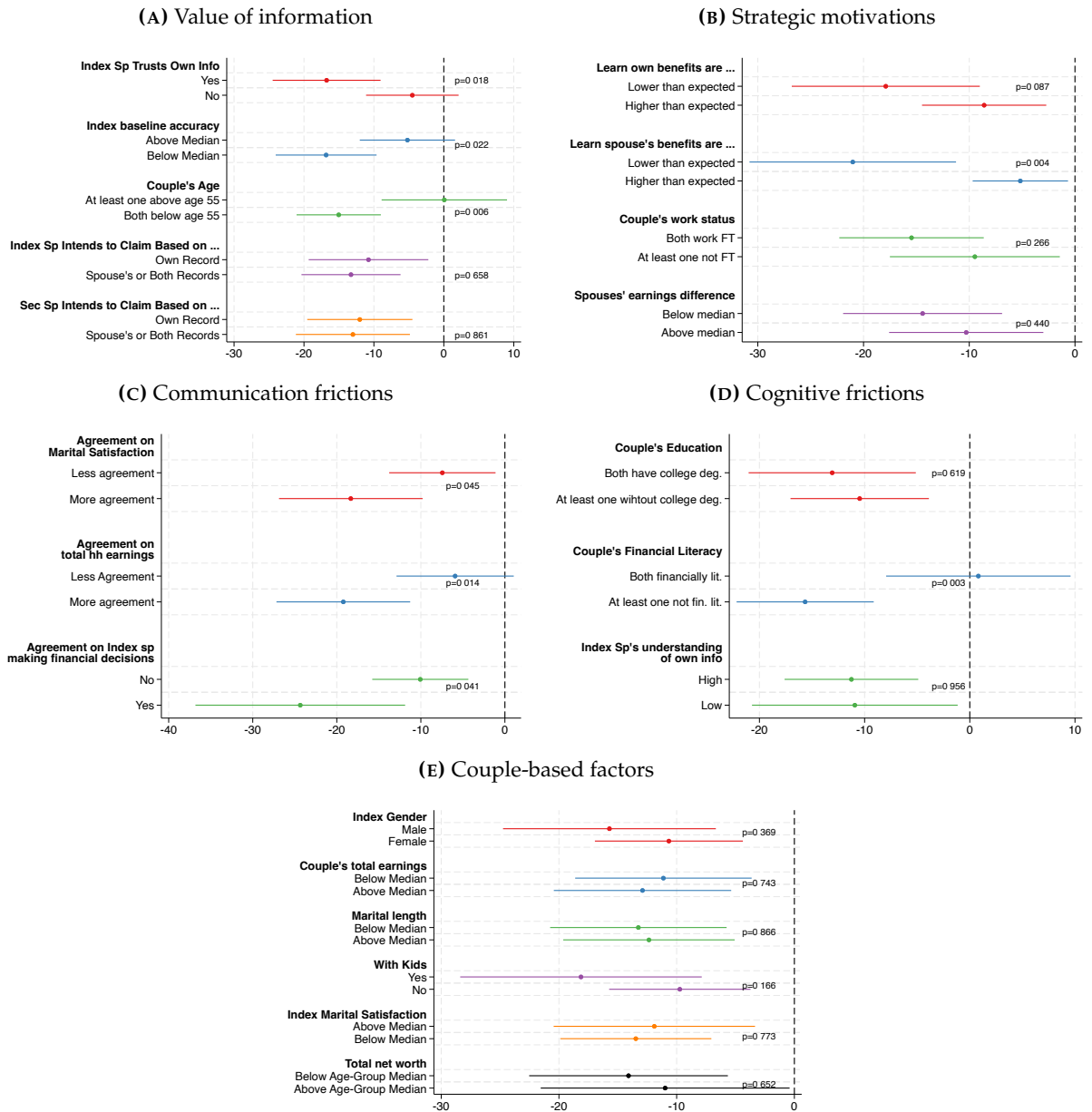
Note: The figure shows the average treatment effects for the index spouse benefits from Stage 0 to Stage 4, estimated using equation (8), separately by different observable characteristics of the index spouse or the couple. The figures include the 90% confidence intervals. Appendix Table A8 presents the estimates for each data cut.

FIGURE A7: Information spillovers to the secondary spouse - With test-retest measurement error



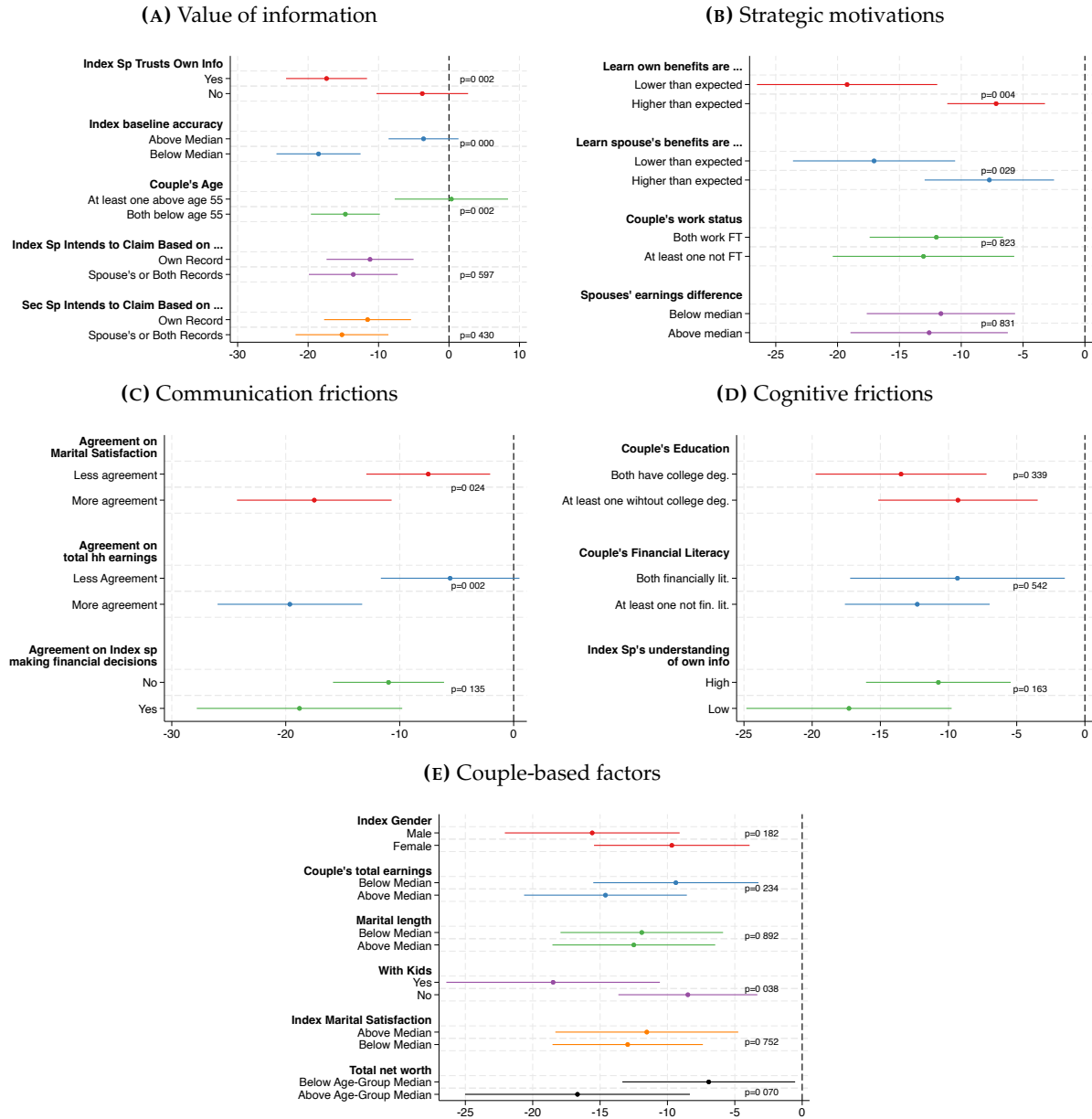
Note: The figure shows the histograms of spillovers in index spouse benefits (left) and secondary spouse benefits (right) as a share of the index spouse's revision in own or spouse's benefits from Stage 0 to Stage 4, across 1000 bootstrap repetitions. For this exercise, we assume that there are full spillovers, so the index spouse's stage 4 beliefs are equal to secondary spouse's stage 5 beliefs, but that secondary spouses report their beliefs with measurement error with a standard deviation equal to the test-retest measurement error. The red vertical dashed line in both figures represent the estimated spillovers as a share of the index spouse treatment effect from the observed data. In the left panel, spillovers with measurement error are higher than the actual spillovers 90.5% of the repetitions, while on the right spillover with measurement error are higher than the actual spillovers 98.7% of the repetitions.

FIGURE A8: Heterogeneity in information spillovers to the secondary spouse - Secondary spouse benefits



Note: Figures show the spillovers in secondary spouse benefits estimated using equation (9) separately for different groups.

FIGURE A9: Heterogeneity in information spillovers about index spouse's benefits to the secondary spouse - Controlling for index spouse's baseline perception gap



Note: Figures show the spillovers in index spouse benefits estimated using equation (9) and controlling for index spouse's baseline perception gap in own benefits, separately for different groups.

TABLE A1: Selection of secondary spouses into completing the survey

| | Pr(sec spouse responding) Control | Pr(sec spouse responding) Treatment | p-val of differences btw C and T |
|---|--------------------------------------|--|-------------------------------------|
| Female | -1.765 (2.03) | -1.061 (1.15) | 0.78 |
| College Grad. | 2.205 (2.35) | 2.635* (1.38) | 0.89 |
| Age above 45 | 0.695 (1.89) | 3.445*** (1.13) | 0.22 |
| Have child under age 6 | -0.919 (2.38) | -1.911 (1.33) | 0.68 |
| Age expected to start claiming SS benefits | -0.121 (0.35) | 0.005 (0.16) | 0.76 |
| Age expected for sp to start claiming SS benefits | 0.117 (0.40) | -0.337* (0.18) | 0.31 |
| Above median earnings | 1.788 (2.33) | -2.607** (1.25) | 0.10 |
| Work full-time | -3.815 (3.25) | -4.597** (1.93) | 0.83 |
| Reports making financial dec. | -0.647 (1.92) | -3.758*** (1.10) | 0.15 |
| Financially literate | 10.228*** (2.53) | 10.538*** (1.39) | 0.94 |
| Above median marital satisf. | -1.608 (1.85) | 0.036 (1.08) | 0.26 |
| Abs perception gap in own benefits (%) | -0.036*** (0.01) | -0.013* (0.01) | 0.10 |
| Abs perception gap in spouse's benefits (%) | 0.030** (0.02) | 0.011 (0.01) | 0.28 |
| Below median marital length | -1.514 (1.89) | 0.440 (1.10) | 0.38 |
| % Chance of Receiving SS Benefits | -0.059 (0.06) | -0.043 (0.03) | 0.80 |
| % Chance of Spouse Receiving SS Benefits | 0.044 (0.06) | 0.066** (0.03) | 0.71 |
| Dep. Var. Mean | 22.568 | 25.522 | |
| R ² | 0.043 | 0.041 | |
| Observations | 2,220 | 6,939 | |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable is a binary variable showing whether the secondary spouse completed the survey. The observable characteristics included in the regression are index spouse's characteristics. All spouse-related variables are reported by the index spouse.

TABLE A2: Comparison of marital satisfaction of spouses within a couple

| Quartiles of Secondary Sp's Marital Satisfaction | Quartiles of Index Sp's Marital Satisfaction | | | |
|---|---|------|------|------|
| | Q1 | Q2 | Q3 | Q4 |
| Q1 | 17.5 | 4.6 | 2.2 | 0.7 |
| Q2 | 6.8 | 11.4 | 5.7 | 1.2 |
| Q3 | 2.5 | 6.9 | 18.3 | 2.9 |
| Q4 | 0.7 | 2.0 | 6.6 | 10.0 |

Note: The table shows within-couple variation in the marital satisfaction index. Each cell in the table shows the share of couples with the index and secondary spouses' marital satisfaction index in the specified quartiles. Note that a higher value of the index indicates greater satisfaction.

TABLE A3: Comparison to the CPS

| | Survey | 2022 CPS |
|--|---------|----------|
| Female | 0.59 | 0.42 |
| White | 0.88 | 0.80 |
| Age < 45 | 0.54 | 0.28 |
| Age 45-50 | 0.19 | 0.26 |
| Age > 50 | 0.27 | 0.45 |
| College degree | 0.65 | 0.47 |
| Employed | 0.85 | 0.86 |
| Weekly earnings (\$) for College Grads | 2033.95 | 1804.50 |
| Weekly earnings(\$) for Non-College | 975.99 | 1098.74 |
| Unemployed | 0.03 | 0.02 |
| Midwest | 0.31 | 0.21 |
| South | 0.37 | 0.38 |
| West | 0.20 | 0.23 |
| Observations | 2,219 | 104,937 |

Note: The table compares the background characteristics of the index spouses from our survey to the married, non-retired household heads in the 2022 CPS. The left-out Census region is Northeast.

TABLE A4: Index Spouse Social Security Expectations at Baseline (Stage 0)

| | Index Spouse | | | |
|---|--------------|---------|---------|---------|
| | All | Control | Treated | P-Value |
| Expected own age 62 benefits (\$) | 2017.6 | 1910.0 | 2048.2 | 0.19 |
| Expected own age 67 benefits (\$) | 2345.9 | 2227.5 | 2379.6 | 0.15 |
| Expected own age 70 benefits (\$) | 2690.9 | 2609.4 | 2714.2 | 0.37 |
| Expected SS benefit claiming age | 66.5 | 66.6 | 66.5 | 0.39 |
| Expected age 62 benefits for the spouse (\$) | 1865.5 | 1839.1 | 1873.1 | 0.69 |
| Expected age 67 benefits for the spouse (\$) | 2270.7 | 2279.3 | 2268.2 | 0.91 |
| Expected age 70 benefits for the spouse (\$) | 2631.8 | 2641.9 | 2628.9 | 0.91 |
| Expected SS benefit claiming age for the spouse | 66.0 | 66.2 | 65.9 | 0.14 |
| Own benefits | | | | |
| Predicted age 62 benefits by the SSA calculator (\$) | 1403.6 | 1385.2 | 1408.9 | 0.51 |
| Absolute gap in own age 62 benefits (\$) | 1131.2 | 1074.7 | 1147.3 | 0.36 |
| Share overestimating | 52% | 51% | 52% | 0.56 |
| Share underestimating | 41% | 42% | 40% | 0.55 |
| Predicted age 67 benefits by the SSA calculator (\$) | 2036.5 | 2010.0 | 2044.1 | 0.51 |
| Absolute gap in own age 67 benefits (\$) | 1209.2 | 1162.5 | 1222.6 | 0.38 |
| Share overestimating | 44% | 44% | 45% | 0.71 |
| Share underestimating | 49% | 49% | 49% | 0.80 |
| Predicted age 70 benefits by the SSA calculator (\$) | 2559.3 | 2525.3 | 2568.9 | 0.50 |
| Absolute gap in own age 70 benefits (\$) | 1409.9 | 1371.7 | 1420.8 | 0.49 |
| Share overestimating | 40% | 40% | 39% | 0.83 |
| Share underestimating | 55% | 54% | 55% | 0.77 |
| Benefits for the spouse | | | | |
| Spouse's predicted age 62 benefits by the SSA calculator (\$) | 1331.9 | 1293.4 | 1342.8 | 0.14 |
| Absolute gap in spouse's age 62 benefits (\$) | 1017.5 | 1005.7 | 1020.8 | 0.80 |
| Share overestimating | 52% | 53% | 52% | 0.59 |
| Share underestimating | 40% | 40% | 41% | 0.73 |
| Spouse's predicted age 67 benefits by the SSA calculator (\$) | 1930.7 | 1876.3 | 1946.2 | 0.15 |
| Absolute gap in spouse's age 67 benefits (\$) | 1210.6 | 1198.0 | 1214.1 | 0.81 |
| Share overestimating | 45% | 47% | 44% | 0.24 |
| Share underestimating | 50% | 47% | 50% | 0.17 |
| Spouse's predicted age 70 benefits by the SSA calculator (\$) | 2431.1 | 2362.9 | 2450.5 | 0.15 |
| Absolute gap in spouse's age 70 benefits (\$) | 1439.0 | 1410.0 | 1447.3 | 0.61 |
| Share overestimating | 40% | 42% | 40% | 0.35 |
| Share underestimating | 54% | 52% | 55% | 0.31 |
| Observations | 2,219 | 492 | 1,727 | |

Note: The table shows the average expected benefits of the index spouses. In the bottom panel of the table, absolute gap is defined as the absolute value of the difference between the predicted benefits by the SS Quick calculator and the benefit expectations of the respondent.

TABLE A5: Heterogeneity in index spouses' absolute perception gaps in own age 67 benefits

| | Average Abs Gap in Own Age 67 benefits | p-value of differences across groups |
|------------------------------|--|--|
| Overall | 53.332*** | |
| Male | 53.350*** | |
| Female | 53.314*** | 0.98 |
| Below age 55 | 54.837*** | |
| Above age 55 | 42.546*** | 0.00 |
| Couple below-median earnings | 49.273*** | |
| Couple above-median earnings | 56.731*** | 0.00 |
| No college | 44.196*** | |
| College grad. | 57.830*** | 0.00 |
| Not financially lit. | 58.428*** | |
| Financially lit. | 43.027*** | 0.00 |

Note: The table shows the average absolute perception gaps of index spouses for their own age 67 benefits as a share of the benefits predicted by the SSA calculator. * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$, show the statistical difference of the average absolute perception gaps from 0. The last column shows the p-value of the differences in the average absolute gaps across the mutually exclusive groups, as defined in the first column.

TABLE A6: Average treatment effects on index spouse's absolute perception gaps in own and spouse's age 67 benefits

| | (1) Rev in Abs PG in own age 67 ben (pp) S0-S2 | (2) Rev in Abs PG in own age 67 ben (pp) S2-S4 | (3) Rev in Abs PG in own age 67 ben (pp) S0-S4 | (4) Rev in Abs PG in sp age 67 ben (pp) S0-S2 | (5) Rev in Abs PG in sp age 67 ben (pp) S2-S4 | (6) Rev in Abs PG in sp age 67 ben (pp) S0-S4 |
|--------------------------|--|--|--|---|---|---|
| Average Treatment Effect | -21.73*** (2.03) | -0.37 (1.34) | -21.98*** (1.90) | -8.98*** (2.01) | -12.99*** (1.46) | -22.69*** (1.98) |
| Baseline Control Abs PG | 55.2 | 54.4 | 54.6 | 55.3 | 55.1 | 54.1 |
| Dep. Var. Mean | -17.46 | 0.08 | -17.51 | -6.53 | -6.95 | -14.04 |
| R ² | 0.06 | 0.01 | 0.06 | 0.02 | 0.05 | 0.06 |
| Observations | 1,507 | 1,438 | 1,535 | 1,558 | 1,466 | 1,553 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in the first column is the revision in the absolute perception gap (PG) of the index spouse in their own age 67 benefits from Stage 0 to Stage 2; in the second column the dependent variable is the revision in absolute PG in index spouse's own age 67 benefits from Stage 2 to 4 and in the third column it is the revision from Stage 0 to 4. The dependent variables in the fourth, fifth and the sixth columns are the equivalent of the dependent variables in the first three columns, for secondary spouse's benefits. The absolute perception gaps in all columns are measured as the absolute perception gaps as shares of benefits calculated by the SSA Quick Calculator. Demographics are also included in all columns and these include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy of both the index and the secondary spouses.

TABLE A7: Treatment effects on index spouse's beliefs about own and spouse's age 67 benefits, by terciles of the baseline perception gap

| | (1) Rev in Exp about own age 67 ben (pp) S0-S2 | (2) Rev in Exp about own age 67 ben (pp) S2-S4 | (3) Rev in Exp about own age 67 ben (pp) S0-S4 | (4) Rev in Exp about sp age 67 ben (pp) S0-S2 | (5) Rev in Exp about sp age 67 ben (pp) S2-S4 | (6) Rev in Exp about sp age 67 ben (pp) S0-S4 |
|--|--|--|--|---|---|---|
| Treatment × PG for Own < p33 | -33.34*** (5.50) | -4.73 (3.40) | -56.41*** (9.23) | | | |
| Treatment × PG for Own ∈ [p33, p66] | 4.88 (3.01) | 2.73 (2.19) | 8.37*** (3.16) | | | |
| Treatment × PG for Own > p66 | 26.21*** (3.59) | 4.06** (1.98) | 30.63*** (4.11) | | | |
| Treatment × PG for Spouse < p33 | | | | -14.44** (6.36) | -22.09*** (4.57) | -46.18*** (12.26) |
| Treatment × PG for Spouse ∈ [p33, p66] | | | | -0.76 (3.25) | -0.94 (2.47) | -1.85 (3.77) |
| Treatment × PG for Spouse > p66 | | | | 19.84*** (4.13) | 8.61*** (2.20) | 30.28*** (4.73) |
| Baseline Control Exp. | 2408.8 | 2380.7 | 2395.6 | 2333.0 | 2331.0 | 2311.2 |
| Dep. Var. Mean | -3.76 | 1.86 | -3.38 | 4.02 | -0.28 | 1.63 |
| R ² | 0.32 | 0.02 | 0.21 | 0.18 | 0.06 | 0.10 |
| Observations | 1,507 | 1,436 | 1,535 | 1,558 | 1,464 | 1,553 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in the first column is the revision in index spouse's beliefs about their own age 67 benefits from Stage 0 to 2; in the second column the dependent variable is the revision in beliefs about own age 67 benefits from Stage 2 to 4 and in the third column it is the revision from Stage 0 to 4. The dependent variables in the fourth, fifth and the sixth columns are the equivalent of the dependent variables in the first three columns, for index spouse's beliefs about the secondary spouse's benefits. The revision in beliefs are measured as percentage point changes. Demographics are also included in all columns and these include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy of both the index and the secondary spouses. All columns also include dummies for the terciles of the baseline perception gaps.

TABLE A8: Heterogeneity in the treatment effects on index spouse's absolute perception gaps about own age 67 benefits

| PANEL A: | | Index Spouse Benefits | | | | | | | | | |
|---------------------------|--------------------------------------|---|--------------------------------------|--|---------------------------|---------------------------|---|--|--------------------------|--------------------------|--|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | |
| | Index Trusts Own Info | Index Doesn't Trust Own Info | Index Sp Above Med Uncertainty | Index Sp Below Med Uncertainty | At least one above 55 | Both sp below 55 | Index Gender Male | Index Gender Female | Total Earn. Below Med | Total Earn. Above Med | |
| Treated | -24.3*** (2.7) | -18.3*** (2.8) | -23.4*** (3.1) | -20.4*** (2.4) | -11.9*** (3.8) | -24.4*** (2.1) | -19.6*** (2.8) | -23.6*** (2.6) | -18.7*** (2.7) | -25.1*** (2.8) | |
| Difference Between Groups | | 6.0 | | 3.1 | | -12.5*** | | -4.0 | | -6.4 | |
| TE as % of CG PG | -39.9 | -39.4 | -38.0 | -41.9 | -29.5 | -42.4 | -34.0 | -45.0 | -37.8 | -42.6 | |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | |
| Control Dep. Var Mean | 0.8 | -1.8 | -0.3 | -0.4 | -6.6 | 0.9 | -1.6 | 0.7 | -0.0 | -0.3 | |
| Adj. R-Squared | 0.0 | 0.1 | 0.0 | 0.1 | 0.0 | 0.1 | 0.0 | 0.1 | 0.0 | 0.1 | |
| Observations | 931 | 604 | 751 | 784 | 270 | 1,265 | 650 | 884 | 687 | 806 | |
| PANEL B: | | Index Spouse Benefits | | | | | | | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | | | |
| | Sp Agree Index Sp Fin. Savvier | Sp Disagree Index Sp Fin. Savvier | Aggr on Index Making Fin. Dec. | Less Aggr on Index Making Fin. Dec. | Less Aggr. on Mar. Sat | More Aggr. on Mar. Sat | Index Learns Own Ben Lower than exp | Index Learns Own Ben Higher than exp | | | |
| Treated | -26.6*** (5.0) | -20.9*** (2.1) | -27.8*** (4.2) | -20.4*** (2.2) | -20.0*** (2.6) | -23.9*** (2.8) | -24.1*** (3.5) | -20.4*** (2.1) | | | |
| Difference Between Groups | | 5.8 | | 7.4 | | -3.9 | | 3.6 | | | |
| TE as % of CG PG | -53.9 | -37.7 | -50.0 | -37.6 | -38.1 | -41.8 | -33.5 | -49.9 | | | |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | | | |
| Control Dep. Var Mean | 2.3 | -0.9 | 0.1 | -0.5 | 0.1 | -0.8 | -2.0 | 1.0 | | | |
| Adj. R-Squared | 0.1 | 0.0 | 0.1 | 0.0 | 0.0 | 0.0 | 0.0 | 0.1 | | | |
| Observations | 220 | 1,304 | 317 | 1,218 | 785 | 749 | 693 | 841 | | | |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in all columns is the revision in index spouse's beliefs about their own age 67 benefits from Stage 0 to 4. The revision in beliefs are measured as percentage point changes. Demographics are also included in all columns and these include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy of both the index and the secondary spouses. In column titles, "Sp Agree Index Sp Fin. Savvier" refers to "Spouses Agree Index Spouse is Financially Savvier" and "Mar. Sat" refers to marital satisfaction.

TABLE A9: Information spillovers to the secondary spouse - Controlling for index spouse's baseline perception gap

| | Percents | | Dollars | |
|-----------------------|----------------------|--------------------|----------------------|---------------------|
| | (1) Index Sp Ben. | (2) Sec Sp Ben. | (3) Index Sp Ben. | (4) Sec Sp Ben. |
| Treated Couple | -12.2*** (2.2) | -10.4*** (2.4) | -304.8*** (57.2) | -236.4*** (62.3) |
| TE as % of CG PG | -22.9 | -18.9 | -25.4 | -19.9 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ |
| Control Dep. Var Mean | 53.3 | 55.1 | 1199.3 | 1189.5 |
| Adj. R-Squared | 0.2 | 0.2 | 0.2 | 0.2 |
| Observations | 2,029 | 2,006 | 2,029 | 2,006 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in the first two columns is the absolute perception gap of the secondary spouse in index spouse's or their own age 67 benefits as share of the predicted benefits by the SS quick calculator. The dependent variable in the last two columns is the the absolute perception gap of the secondary spouse in index spouse's or their own age 67 benefits in dollars. All four regressions also control for the index spouse's baseline perception gap in their own (columns 1 and 3) or their spouse's (columns 2 and 4) benefits. Demographics include the ethnicity, education, gender, age group, and numeracy of both the index and the secondary spouses.

TABLE A10: Information spillovers to the secondary spouse by value of information

| PANEL A: | | | | | | | | | | |
|--|-----------------------|------------------------------|---------------------------------|---------------------------------|-----------------------|-------------------|---------------------------------|--|----------------------------------|--|
| Index Spouse Benefits | | | | | | | | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| | Index Trusts Own Info | Index Doesn't Trust Own Info | Exp. Ben. Accuracy Above Median | Exp. Ben. Accuracy Below Median | At least one above 55 | Both sp below 55 | Index Claim Based on Own Record | Index Claim Based on Sp or Both Record | Sec Sp Claim Based on Own Record | Sec Sp Claim Based on Sp or Both Records |
| Treated | -17.8*** (3.3) | -3.8 (3.4) | -3.5 (2.5) | -18.6*** (3.5) | 1.7 (4.4) | -15.7*** (2.7) | -11.1*** (3.6) | -13.2*** (3.5) | -10.1*** (3.4) | -17.8*** (3.7) |
| Difference Between Groups | | 13.9*** | | -15.1*** | | -17.4*** | | -2.1 | | -7.6 |
| Control Dep. Var Mean | 58.2 | 47.4 | 33.5 | 72.6 | 34.6 | 57.2 | 50.0 | 55.5 | 51.0 | 57.2 |
| PANEL B: | | | | | | | | | | |
| Index Spouse Benefits, controlling for baseline PG | | | | | | | | | | |
| Treated | -17.4*** (2.9) | -3.8 (3.3) | -3.6 (2.5) | -18.5*** (3.0) | 0.3 (4.2) | -14.7*** (2.5) | -11.2*** (3.1) | -13.6*** (3.2) | -11.6*** (3.1) | -15.2*** (3.4) |
| Difference Between Groups | | 13.6*** | | -14.9*** | | -15.1*** | | -2.4 | | -3.6 |
| Control Dep. Var Mean | 58.2 | 47.4 | 33.5 | 72.6 | 34.6 | 57.2 | 50.0 | 55.5 | 51.0 | 57.2 |
| PANEL C: | | | | | | | | | | |
| Index Spouse Benefits, normalized using z-scores | | | | | | | | | | |
| Treated | -0.4*** (0.1) | -0.1 (0.1) | -0.1 (0.1) | -0.4*** (0.1) | 0.1 (0.1) | -0.3*** (0.1) | -0.2*** (0.1) | -0.3*** (0.1) | -0.2*** (0.1) | -0.4*** (0.1) |
| Difference Between Groups | | 0.3** | | -0.3** | | -0.4*** | | -0.0 | | -0.2 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Control Dep. Var Mean | 0.2 | 0.1 | 0.1 | 0.2 | -0.0 | 0.2 | 0.2 | 0.2 | 0.1 | 0.2 |
| Observations | 1,214 | 815 | 1,010 | 1,019 | 376 | 1,653 | 903 | 972 | 1,005 | 867 |
| PANEL D: | | | | | | | | | | |
| Secondary Spouse Benefits | | | | | | | | | | |
| Treated | -16.8*** (3.9) | -4.5 (3.4) | -5.2 (3.5) | -16.8*** (3.7) | 0.1 (4.6) | -15.0*** (3.1) | -10.8** (4.4) | -13.3*** (3.6) | -12.0*** (3.8) | -13.0*** (4.2) |
| Difference Between Groups | | 12.3** | | -11.6** | | -15.1*** | | -2.5 | | -1.0 |
| Control Dep. Var Mean | 63.3 | 45.7 | 40.0 | 68.6 | 35.1 | 59.5 | 54.4 | 55.3 | 50.0 | 61.0 |
| PANEL E: | | | | | | | | | | |
| Secondary Spouse Benefits, controlling for baseline PG | | | | | | | | | | |
| Treated | -15.3*** (3.4) | -3.4 (3.3) | -3.3 (3.2) | -15.0*** (3.3) | -0.8 (4.5) | -12.1*** (2.8) | -9.8*** (3.8) | -11.1*** (3.4) | -10.5*** (3.4) | -10.2*** (3.9) |
| Difference Between Groups | | 11.9** | | -11.6** | | -11.2** | | -1.4 | | 0.2 |
| Control Dep. Var Mean | 63.3 | 45.7 | 40.0 | 68.6 | 35.1 | 59.5 | 54.4 | 55.3 | 50.0 | 61.0 |
| PANEL F: | | | | | | | | | | |
| Secondary Spouse Benefits, normalized using z-scores | | | | | | | | | | |
| Treated | -0.3*** (0.1) | -0.1 (0.1) | -0.1 (0.1) | -0.3*** (0.1) | 0.0 (0.1) | -0.3*** (0.1) | -0.2** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) |
| Difference Between Groups | | 0.2* | | -0.2* | | -0.3** | | -0.1 | | 0.0 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Control Dep. Var Mean | 0.2 | 0.1 | 0.1 | 0.2 | 0.0 | 0.2 | 0.1 | 0.1 | 0.2 | 0.1 |
| Observations | 1,165 | 841 | 964 | 1,042 | 380 | 1,626 | 853 | 986 | 1,010 | 835 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in all columns is the absolute baseline perception gap as a share of predicted age 67 benefits by the SS quick calculator. Robust standard errors are in parentheses. Demographics include the ethnicity, education, gender, age group, and numeracy for both the index and the secondary spouses. Panels A-C show the spillovers in index spouse's benefits and D-F show the spillovers in secondary spouse's benefits. Panels A and D show estimates from equation (9), panels B and E show estimates from equation (9) also controlling for the baseline perception gap of the index spouse and panels C and F show the estimates from equation (9) by using within-group normalized (using a z-score) perception gaps as the dependent variable. In column titles, "Sec sp" refers to secondary spouse, "Age Gr Med" refers to Age Group Median.

TABLE A11: Information spillovers to the secondary spouse by strategic motivation

| PANEL A: | | Index Spouse Benefits | | | | | | |
|---------------------------|--|--|---------------------------------------|--|---------------------------|---------------------------------------|-------------------------------------|-------------------------------------|
| | (1) Learn Own Ben Lower than exp | (2) Learn Own Ben Higher than exp | (3) Learn Sp Ben Lower than exp | (4) Learn Sp Ben Higher than exp | (5) Both sp work FT | (6) At least one sp not work FT | (7) Sp Earn Diff Below Median | (8) Sp Earn Diff Above Median |
| Treated | -22.1*** (4.6) | -7.0*** (2.2) | -18.8*** (4.0) | -7.8*** (2.6) | -13.6*** (3.0) | -12.9*** (4.0) | -12.5*** (3.3) | -13.0*** (3.6) |
| Difference Between Groups | | 15.1*** | | 11.0** | | 0.7 | | -0.5 |
| Control Dep. Var Mean | 67.2 | 43.5 | 61.5 | 46.0 | 53.6 | 52.7 | 53.2 | 53.1 |
| PANEL B: | | Index Spouse Benefits, controlling for baseline PG | | | | | | |
| Treated | -19.2*** (3.7) | -7.2*** (2.0) | -17.1*** (3.3) | -7.7*** (2.7) | -12.0*** (2.7) | -13.1*** (3.8) | -11.6*** (3.1) | -12.6*** (3.2) |
| Difference Between Groups | | 12.1*** | | 9.3** | | -1.0 | | -0.9 |
| Control Dep. Var Mean | 67.2 | 43.5 | 61.5 | 46.0 | 53.6 | 52.7 | 53.2 | 53.1 |
| PANEL C: | | Index Spouse Benefits, normalized using z-scores | | | | | | |
| Treated | -0.398*** (0.1) | -0.225*** (0.1) | -0.356*** (0.1) | -0.228*** (0.1) | -0.296*** (0.1) | -0.313*** (0.1) | -0.275*** (0.1) | -0.303*** (0.1) |
| Difference Between Groups | | 0.174 | | 0.128 | | -0.017 | | -0.027 |
| Control Dep. Var Mean | 0.2 | 0.2 | 0.2 | 0.2 | 0.1 | 0.2 | 0.1 | 0.2 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Observations | 875 | 1,149 | 928 | 1,085 | 1,220 | 809 | 973 | 1,007 |
| PANEL D: | | Secondary Spouse Benefits | | | | | | |
| Treated | -17.9*** (4.5) | -8.6*** (3.0) | -21.0*** (5.0) | -5.2** (2.3) | -15.5*** (3.5) | -9.5** (4.1) | -14.4*** (3.8) | -10.3*** (3.7) |
| Difference Between Groups | | 9.3* | | 15.9*** | | 6.0 | | 4.1 |
| Control Dep. Var Mean | 66.0 | 46.8 | 72.2 | 41.8 | 57.2 | 51.7 | 55.7 | 54.0 |
| PANEL E: | | Secondary Spouse Benefits, controlling for baseline PG | | | | | | |
| Treated | -14.2*** (4.0) | -8.4*** (3.0) | -17.6*** (4.2) | -6.1*** (2.1) | -11.2*** (3.1) | -9.5** (3.9) | -11.7*** (3.4) | -9.0*** (3.5) |
| Difference Between Groups | | 5.8 | | 11.5** | | 1.8 | | 2.7 |
| Control Dep. Var Mean | 66.0 | 46.8 | 72.2 | 41.8 | 57.2 | 51.7 | 55.7 | 54.0 |
| PANEL F: | | Secondary Spouse Benefits, normalized using z-scores | | | | | | |
| Treated | -0.316*** (0.1) | -0.228*** (0.1) | -0.351*** (0.1) | -0.156** (0.1) | -0.309*** (0.1) | -0.212** (0.1) | -0.295*** (0.1) | -0.217*** (0.1) |
| Difference Between Groups | | 0.088 | | 0.195* | | 0.097 | | 0.078 |
| Control Dep. Var Mean | 0.2 | 0.2 | 0.2 | 0.1 | 0.2 | 0.1 | 0.2 | 0.1 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Observations | 897 | 1,097 | 860 | 1,143 | 1,202 | 804 | 944 | 1,020 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in all columns is the absolute baseline perception gap as a share of predicted age 67 benefits by the SS quick calculator. Robust standard errors are in parentheses. Demographics include the ethnicity, education, gender, age group, and numeracy for both the index and the secondary spouses. Panels A-C show the spillovers in index spouse's benefits and D-F show the spillovers in secondary spouse's benefits. Panels A and D show estimates from equation (9), panels B and E show estimates from equation (9) also controlling for the baseline perception gap of the index spouse and panels C and F show the estimates from equation (9) by using within-group normalized (using a z-score) perception gap as the dependent variable.

TABLE A12: Information spillovers to the secondary spouse by communication factors

| PANEL A: | | | | | | |
|--|----------------------------------|----------------------------------|-----------------------------------|-----------------------------------|---|--|
| Index Spouse Benefits | | | | | | |
| | (1) Less Aggr. on Mar. Sat | (2) More Aggr. on Mar. Sat | (3) Less Aggr on Total Earn | (4) More Aggr on Total Earn | (5) Less Aggr on Index Making Fin. Dec. | (6) Aggr on Index Making Fin. Dec. |
| Treated | -7.4** (2.9) | -19.0*** (3.9) | -6.6** (3.3) | -20.1*** (3.5) | -11.6*** (2.7) | -19.3*** (5.0) |
| Difference Between Groups | | -11.7** | | -13.5*** | | -7.7 |
| Control Dep. Var Mean | 50.2 | 56.9 | 51.7 | 54.6 | 52.9 | 54.7 |
| PANEL B: | | | | | | |
| Index Spouse Benefits, controlling for baseline PG | | | | | | |
| Treated | -7.5*** (2.8) | -17.5*** (3.5) | -5.6* (3.1) | -19.6*** (3.2) | -11.0*** (2.5) | -18.8*** (4.7) |
| Difference Between Groups | | -10.2** | | -14.0*** | | -7.6 |
| Control Dep. Var Mean | 50.2 | 56.9 | 51.7 | 54.6 | 52.9 | 54.7 |
| PANEL C: | | | | | | |
| Index Spouse Benefits, normalized using z-scores | | | | | | |
| Treated | -0.178** (0.1) | -0.405*** (0.1) | -0.146** (0.1) | -0.480*** (0.1) | -0.257*** (0.1) | -0.485*** (0.1) |
| Difference Between Groups | | -0.228** | | -0.334*** | | -0.228 |
| Control Dep. Var Mean | 0.1 | 0.2 | 0.1 | 0.3 | 0.1 | 0.4 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Observations | 1,026 | 1,002 | 966 | 946 | 1,627 | 402 |
| PANEL D: | | | | | | |
| Secondary Spouse Benefits | | | | | | |
| Treated | -7.4** (3.2) | -18.3*** (4.4) | -5.9* (3.6) | -19.2*** (4.1) | -10.1*** (2.9) | -24.3*** (6.5) |
| Difference Between Groups | | -10.9** | | -13.3** | | -14.3** |
| Control Dep. Var Mean | 51.5 | 59.4 | 53.1 | 56.9 | 52.1 | 66.7 |
| PANEL E: | | | | | | |
| Secondary Spouse Benefits, controlling for baseline PG | | | | | | |
| Treated | -6.0** (3.0) | -15.2*** (3.9) | -3.1 (3.4) | -17.6*** (3.6) | -7.3*** (2.6) | -23.9*** (5.9) |
| Difference Between Groups | | -9.4* | | -14.5*** | | -16.6*** |
| Control Dep. Var Mean | 51.5 | 59.4 | 53.1 | 56.9 | 52.1 | 66.7 |
| PANEL F: | | | | | | |
| Secondary Spouse Benefits, normalized using z-scores | | | | | | |
| Treated | -0.166** (0.1) | -0.359*** (0.1) | -0.120* (0.1) | -0.414*** (0.1) | -0.209*** (0.1) | -0.513*** (0.1) |
| Difference Between Groups | | -0.192* | | -0.294*** | | -0.304** |
| Control Dep. Var Mean | 0.1 | 0.2 | 0.0 | 0.3 | 0.1 | 0.4 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Observations | 1,027 | 978 | 982 | 957 | 1,627 | 379 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in all columns is the absolute baseline perception gap as a share of predicted age 67 benefits by the SS quick calculator. Robust standard errors are in parentheses. Demographics include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy of both the index and the secondary spouses. Panels A-C show the spillovers in index spouse's benefits and D-F show the spillovers in secondary spouse's benefits. Panels A and D show estimates from equation (9), panels B and E show estimates from equation (9) also controlling for the baseline perception gap of the index spouse and panels C and F show the estimates from equation (9) by using within-group normalized (using a z-score) perception gaps as the dependent variable.

TABLE A13: Information spillovers to the secondary spouse by cognitive factors

| PANEL A: Index Spouse Benefits | | | | | | |
|---|---------------------|---------------------------------|----------------------|----------------------------------|---------------------------------------|--------------------------------------|
| | (1) Both college | (2) At least one not college | (3) Both Fin. Lit | (4) At least one Not Fin. Lit | (5) Index unders. check high score | (6) Index unders. check low score |
| Treated | -14.4*** (3.6) | -9.2*** (3.1) | -8.8** (4.1) | -13.1*** (3.0) | -12.3*** (2.9) | -13.0*** (4.3) |
| Difference Between Groups | | 5.2 | | -4.3 | | -0.8 |
| Control Dep. Var Mean | 58.6 | 47.6 | 45.9 | 55.4 | 51.3 | 56.7 |
| PANEL B: Index Spouse Benefits, controlling for baseline PG | | | | | | |
| Treated | -13.5*** (3.2) | -9.3*** (3.0) | -9.3** (4.1) | -12.3*** (2.7) | -10.7*** (2.7) | -17.3*** (3.9) |
| Difference Between Groups | | 4.2 | | -3.0 | | -6.6 |
| Control Dep. Var Mean | 58.6 | 47.6 | 45.9 | 55.4 | 51.3 | 56.7 |
| PANEL C: Index Spouse Benefits, normalized using z-scores | | | | | | |
| Treated | -0.3*** (0.1) | -0.3*** (0.1) | -0.2** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) |
| Difference Between Groups | | 0.0 | | -0.1 | | 0.0 |
| Control Dep. Var Mean | 0.2 | 0.2 | 0.2 | 0.2 | 0.2 | 0.1 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Observations | 1,117 | 912 | 487 | 1,542 | 1,446 | 583 |
| PANEL D: Secondary Spouse Benefits | | | | | | |
| Treated | -13.1*** (4.0) | -10.5*** (3.4) | 0.8 (4.5) | -15.7*** (3.3) | -11.3*** (3.2) | -10.9** (5.0) |
| Difference Between Groups | | 2.6 | | -16.5*** | | 0.3 |
| Control Dep. Var Mean | 61.5 | 48.7 | 39.7 | 59.5 | 51.1 | 61.6 |
| PANEL E: Secondary Spouse Benefits, controlling for baseline PG | | | | | | |
| Treated | -9.6*** (3.5) | -10.1*** (3.2) | 0.9 (4.5) | -12.9*** (3.0) | -9.6*** (3.0) | -11.4** (4.4) |
| Difference Between Groups | | -0.5 | | -13.8** | | -1.8 |
| Control Dep. Var Mean | 61.5 | 48.7 | 39.7 | 59.5 | 51.1 | 61.6 |
| PANEL F: Secondary Spouse Benefits, normalized using z-scores | | | | | | |
| Treated | -0.2*** (0.1) | -0.3*** (0.1) | 0.0 (0.1) | -0.3*** (0.1) | -0.3*** (0.1) | -0.2** (0.1) |
| Difference Between Groups | | -0.0 | | -0.3*** | | 0.0 |
| Control Dep. Var Mean | 0.1 | 0.2 | 0.0 | 0.2 | 0.2 | 0.1 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Observations | 1,080 | 926 | 487 | 1,519 | 1,440 | 566 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in all columns is the absolute baseline perception gap as a share of predicted age 67 benefits by the SS quick calculator. Robust standard errors are in parentheses. Demographics include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy of both the index and the secondary spouses. Panels A-C show the spillovers in index spouse's benefits and D-F show the spillovers in secondary spouse's benefits. Panels A and D show estimates from equation (9), panels B and E show estimates from equation (9) also controlling for the baseline perception gap of the index spouse and panels C and F show the estimates from equation (9) by using within-group normalized (using a z-score) perception gaps as the dependent variable.

TABLE A14: Information spillovers to the secondary spouse by couple-based factors

| PANEL A: | | | | | | | | | | | | |
|--|-----------------------------|-------------------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|------------------------------|----------------------------|-----------------------------------|------------------------------------|---|---|
| Index Spouse Benefits | | | | | | | | | | | | |
| | (1) Index Gender Male | (2) Index Gender Female | (3) Total Earn. Below Med | (4) Total Earn. Above Med | (5) Below Med Mar. Length | (6) Above Med Mar. Length | (7) With Child Under18 | (8) No Child Under18 | (9) Index Mar Sat Above Med | (10) Index Mar Sat Below Med | (11) Total net worth Below Age Gr Med | (12) Total net worth Above Age Gr Med |
| Treated | -15.8*** (3.7) | -10.0*** (3.2) | -10.2*** (3.3) | -15.1*** (3.4) | -13.7*** (3.4) | -12.3*** (3.3) | -19.5*** (4.5) | -9.1*** (2.8) | -13.1*** (3.9) | -12.8*** (3.0) | -7.3** (3.6) | -15.8*** (4.7) |
| Difference Between Groups | | 5.8 | | -4.9 | | 1.4 | | 10.4* | | 0.3 | | -8.5 |
| Control Dep. Var Mean | 57.1 | 50.5 | 49.0 | 56.6 | 53.5 | 53.1 | 61.3 | 49.7 | 57.3 | 50.1 | 49.4 | 58.9 |
| PANEL B: | | | | | | | | | | | | |
| Index Spouse Benefits, controlling for baseline PG | | | | | | | | | | | | |
| Treated | -15.6*** (3.3) | -9.7*** (2.9) | -9.4*** (3.1) | -14.6*** (3.1) | -11.9*** (3.1) | -12.5*** (3.1) | -18.5*** (4.1) | -8.5*** (2.6) | -11.5*** (3.5) | -13.0*** (2.8) | -6.9** (3.3) | -16.7*** (4.3) |
| Difference Between Groups | | 5.9 | | -5.2 | | -0.6 | | 10.0** | | -1.4 | | -9.7* |
| Control Dep. Var Mean | 57.1 | 50.5 | 49.0 | 56.6 | 53.5 | 53.1 | 61.3 | 49.7 | 57.3 | 50.1 | 49.4 | 58.9 |
| PANEL C: | | | | | | | | | | | | |
| Index Spouse Benefits, normalized using z-scores | | | | | | | | | | | | |
| Treated | -0.4*** (0.1) | -0.2*** (0.1) | -0.2*** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) | -0.4*** (0.1) | -0.2*** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) | -0.2** (0.1) | -0.3*** (0.1) |
| Difference Between Groups | | 0.1 | | -0.1 | | 0.0 | | 0.2* | | -0.1 | | -0.2 |
| Control Dep. Var Mean | 0.3 | 0.1 | 0.1 | 0.2 | 0.1 | 0.2 | 0.3 | 0.1 | 0.2 | 0.2 | 0.1 | 0.2 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Observations | 863 | 1,163 | 924 | 1,056 | 948 | 1,081 | 687 | 1,342 | 977 | 1,052 | 691 | 715 |
| PANEL D: | | | | | | | | | | | | |
| Secondary Spouse Benefits | | | | | | | | | | | | |
| Treated | -15.7*** (4.6) | -10.7*** (3.2) | -11.1*** (3.8) | -12.9*** (3.8) | -13.3*** (3.8) | -12.4*** (3.7) | -18.1*** (5.3) | -9.7*** (3.1) | -11.9*** (4.4) | -13.5*** (3.3) | -14.1*** (4.3) | -11.0** (5.4) |
| Difference Between Groups | | 5.0 | | -1.8 | | 0.9 | | 8.4 | | -1.6 | | 3.1 |
| Control Dep. Var Mean | 62.9 | 50.2 | 51.1 | 58.1 | 56.3 | 54.1 | 63.4 | 51.6 | 57.8 | 53.1 | 57.5 | 60.0 |
| PANEL E: | | | | | | | | | | | | |
| Secondary Spouse Benefits, controlling for baseline PG | | | | | | | | | | | | |
| Treated | -13.9*** (4.2) | -8.6*** (2.9) | -8.0** (3.6) | -12.0*** (3.4) | -10.6*** (3.5) | -10.6*** (3.4) | -13.1*** (4.7) | -8.6*** (2.8) | -8.6** (3.9) | -12.4*** (3.1) | -10.5*** (3.9) | -12.9*** (4.7) |
| Difference Between Groups | | 5.3 | | -4.0 | | 0.0 | | 4.6 | | -3.8 | | -2.3 |
| Control Dep. Var Mean | 62.9 | 50.2 | 51.1 | 58.1 | 56.3 | 54.1 | 63.4 | 51.6 | 57.8 | 53.1 | 57.5 | 60.0 |
| PANEL F: | | | | | | | | | | | | |
| Secondary Spouse Benefits, normalized using z-scores | | | | | | | | | | | | |
| Treated | -0.3*** (0.1) | -0.2*** (0.1) | -0.3*** (0.1) | -0.2*** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) | -0.4*** (0.1) | -0.2*** (0.1) | -0.2*** (0.1) | -0.3*** (0.1) | -0.3*** (0.1) | -0.2** (0.1) |
| Difference Between Groups | | 0.1 | | 0.0 | | 0.0 | | 0.1 | | -0.1 | | 0.1 |
| Control Dep. Var Mean | 0.2 | 0.1 | 0.2 | 0.1 | 0.1 | 0.2 | 0.2 | 0.1 | 0.1 | 0.2 | 0.2 | 0.1 |
| Demog. Controls | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ | ✓ |
| Observations | 797 | 1,206 | 942 | 1,022 | 934 | 1,072 | 659 | 1,347 | 947 | 1,059 | 684 | 680 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in all columns is the absolute baseline perception gap as a share of predicted age 67 benefits by the SS quick calculator. Robust standard errors are in parentheses. Demographics include the ethnicity, education, gender, age group, perception of own financial savviness, and numeracy of both the index and the secondary spouses. Panels A-C show the spillovers in index spouse's benefits and D-F show the spillovers in secondary spouse's benefits. Panels A and D show estimates from equation (9), panels B and E show estimates from equation (9) also controlling for the baseline perception gap of the index spouse and panels C and F show the estimates from equation (9) by using within-group normalized (using a z-score) perception gaps as the dependent variable.

TABLE A15: Spillovers to the secondary spouse with all indices and index spouse's baseline perception gaps

| | Index Sp Benefits | | Sec Sp Benefits | |
|----------------------------------|-------------------|-------------------|------------------|------------------|
| | (1) | (2) | (3) | (4) |
| Treated | -12.1*** (2.4) | -16.0*** (5.7) | -9.8*** (2.6) | -12.9** (6.0) |
| Treated × Lack of cogn barriers | 1.3 (1.9) | 1.7 (2.1) | 3.4 (2.3) | 3.5 (2.7) |
| Treated × Lack of comm frictions | -7.4*** (2.2) | -7.6*** (2.3) | -9.4*** (2.4) | -9.8*** (2.5) |
| Treated × High value of info | -4.8** (2.2) | -4.9** (2.2) | -4.1* (2.5) | -4.2 (2.5) |
| Treated × Negative news | -5.5** (2.2) | -5.6** (2.3) | -6.7*** (2.6) | -7.6*** (2.6) |
| Treated × Equal couples | 1.9 (2.2) | 1.3 (2.2) | 0.9 (2.3) | 0.5 (2.4) |
| Couple-based Factors | | ✓ | | ✓ |
| Control Dep. Var Mean | 53.3 | 53.3 | 55.1 | 55.1 |
| Adj. R-Squared | 0.2 | 0.2 | 0.2 | 0.2 |
| Observations | 2,029 | 2,029 | 2,006 | 2,006 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in all columns is the absolute baseline perception gap as a share of predicted age 67 benefits by the SS quick calculator. Robust standard errors are in parentheses. The specifications also include the uninteracted indices. The regression also controls for the index spouse's baseline perception gaps in their beliefs about own (columns (1) and (2)) or in spouse's (columns (3) and (4)) benefits. Demographics include the ethnicity and perception of own financial savviness of both the index and the secondary spouses.

TABLE A16: Average treatment effects on index spouse's expectations for their own and spouses' claiming age

| | (1) Rev in exp for own cl age S0-S4 | (2) Rev in exp for sp cl age S0-S4 |
|--------------------------|---|--|
| Average Treatment Effect | 0.35** (0.14) | 0.46*** (0.14) |
| Baseline Control Cl Age | 66.6 | 66.2 |
| Dep. Var. Mean | 0.33 | 0.34 |
| R ² | 0.02 | 0.02 |
| Observations | 2,215 | 2,204 |

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. The dependent variable in the first column is the revision in the expected claiming age of the index spouse in their own age 67 benefits from Stage 0 to Stage 2; in the second column the dependent variable is the revision in index spouse's expected own claiming age from Stage 2 to 4 and in the third column it is the revision from Stage 0 to 4. The dependent variables in the fourth, fifth and the sixth columns are the equivalent of the dependent variables in the first three columns, for secondary spouse's claiming age. Demographics are also included in all columns and these include the ethnicity, education, gender, age group, financial literacy and numeracy of both the index and the secondary spouses.

Appendix B Survey Instrument

FIGURE B1: Experimental design

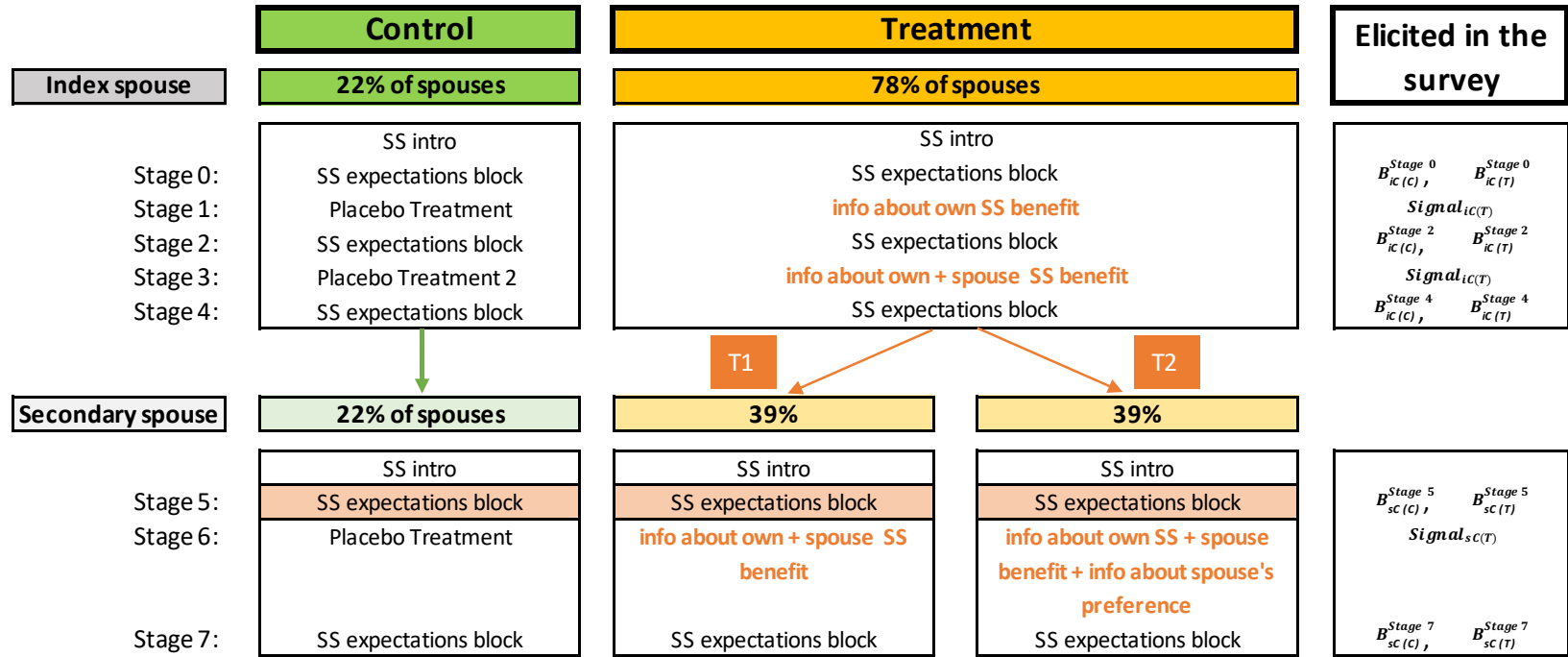


FIGURE B2: Eliciting expectations about own Social Security benefits

If you were to retire at [age 62](#) and 1 month and collect benefits, on average, how much do you expect the monthly payments to be in today's dollars?
Please enter a number in the box below.

dollars per month

If you were to retire at [age 67](#) and collect benefits, on average, how much do you expect the monthly payments to be in today's dollars?
Please enter a number in the box below.

dollars per month

If you were to retire at [age 70](#) and collect benefits, on average, how much do you expect the monthly payments to be in today's dollars?
Please enter a number in the box below.

dollars per month

FIGURE B3: Eliciting expectations about apouse's Social Security benefits

If [your spouse](#) were to retire at [age 62](#) and 1 month and collect benefits, on average, how much do you expect those monthly payments to be in today's dollars?
Please enter a number in the box below.

dollars per month

If [your spouse](#) were to retire at [age 67](#) and collect benefits, on average, how much do you expect those monthly payments to be in today's dollars?
Please enter a number in the box below.

dollars per month

If [your spouse](#) were to retire at [age 70](#) and collect benefits, on average, how much do you expect those monthly payments to be in today's dollars?
Please enter a number in the box below.

dollars per month

FIGURE B4: Information treatments on Social Security benefits

(A) Stage 1 Treatment Information

Based on a formula similar to the one used by the [Social Security Quick Calculator](#) and the earnings you reported earlier, we present below your estimated Social Security benefit amounts for 3 different claiming ages. For these calculations, we assume you will work every year up to the year in which you begin receiving benefits.

Social Security Benefit Estimates

| Retirement and Social Security claiming age | Monthly benefit amount ¹ |
|---|-------------------------------------|
| 62 and 1 month | \$2100 |
| 67 | \$3065 |
| 70 | \$3868 |

As shown in the table, you can receive a monthly benefit starting at age 62 and 1 month that would be reduced for life due to early retirement. If you choose to delay the start of benefits to a higher age, you can then receive a larger monthly benefit for the rest of your life.

For example, if you start taking benefits at age 62 and 1 month, you will receive \$2100 per month for the rest of your life. But if you wait until age 67 to start receiving benefits, you will get \$3065 for the rest of your life. So by waiting until age 67, you can then receive \$965 more per month than if you started claiming at 62 and 1 month. Note that if you start claiming your benefits after age 70, your monthly benefits will still be \$3868. Remember, these estimated figures are in today's dollars.

¹For this estimate, we also assumed no future increases in prices or earnings. We have calculated your benefits based on the earnings you have reported earlier and by making certain assumptions about your past earnings.

(B) Stage 3 Treatment Information

Based on a formula similar to the one used by the [Social Security Quick Calculator](#) and the earnings you reported for your spouse, we present below the estimated Social Security benefit amounts for your spouse for 3 different claiming ages. For these calculations, we assume your spouse will work every year up to the year in which your spouse begins receiving benefits.

Your Spouse's Social Security Benefit Estimates

| Spouse's retirement and Social Security claiming age | Spouse's monthly benefit amount ¹ |
|--|--|
| 62 and 1 month | \$1995 |
| 67 | \$2910 |
| 70 | \$3671 |

Note that if your spouse starts claiming your benefits after age 70, his/her monthly benefits will still be \$3671. For your convenience, we have also reproduced below the table presenting the estimates for your own Social Security estimates that you saw earlier.

Your Social Security Benefit Estimates

| Retirement and Social Security claiming age | Monthly benefit amount ¹ |
|---|-------------------------------------|
| 62 and 1 month | \$2100 |
| 67 | \$3065 |
| 70 | \$3868 |

¹For this estimate, we also assumed no future increases in prices or earnings. We have calculated the benefits based on the earnings you have reported earlier for you and your spouse and by making certain assumptions about your past earnings.

FIGURE B5: Placebo treatments on Social Security benefits

(A) Stage 1 Control Information



Although possibly the most widely played and watched sport globally, men's soccer has long been dominated by South America and Europe. These continents include all previous World Cup winners and a majority of the previous World Cup hosts. Only two countries outside of these continents have ever made it to the semi-finals. At the first World Cup in 1930, the United States took third place, its best finish to date, and in 2002 co-host South Korea finished in fourth place.

Despite the concentration of powerhouses in South America and Europe, soccer hasn't lagged elsewhere. In the United States, Major League Soccer viewership continues to climb compared to the steady NBA and declining NFL viewership, and despite the Premier League's home in England, Asia accounts for almost 40% of its global viewers. Online, top players Cristiano Ronaldo and Lionel Messi have the two most-followed accounts on Instagram held by individuals, further demonstrating the sport's global fandom.

That global support may help distribute the growth of emerging talents, which could be helped by the next two World Cups. In 2022, Qatar will host the tournament in the Middle East, and in 2026, the United States, Canada, and Mexico will co-host in North America. With so much support and interest around the world, soccer shouldn't have to wait long before it sees a new World Cup winner.

The winners of the previous men's Soccer World Cups have been from which of the following continents?
Please select all that apply.

- ☐ North America
- ☐ South America
- ☐ Europe
- ☐ Africa
- ☐ Asia
- ☐ Australia

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(B) Stage 3 Control Information



Swimming has been part of the Olympic schedule since the very first modern Olympic Games in 1896. It's one of only four disciplines to have been retained, appearing in every summer Olympics since – the others being athletics, artistic gymnastics and fencing.

Swimming at the Olympics has changed a lot since 1896, where there were just four men's races, all held in the Mediterranean Sea. It wasn't until the London Games in 1908 when swimming at the Olympics was contested in a pool, with organizers building a 100m long pool in the middle of the athletics running track. Women only started competing in the Olympics swimming competitions in 1912 – 16 years after men.

The post-World War II era brought better technology, facilities and improved training techniques, resulting in significantly quicker times compared to the early, wave-fighting competitions. Originally, female and male swimmers wore body suits, which increased resistance and slowed them down. As the sport progressed, swimwear became more hydrodynamic. Suits began to be made from materials such as Lycra, which reduced drag and, as a result, reduced lap times.

Competitive pools also saw great change during this period, which led to the move from outdoor to indoor tournaments. The introduction of drainage in Olympic swimming pools, marked lanes in 1924, and guidelines for pool depths all contributed to a better overall standard of competition in the years that followed.

Based on this information, in which Olympics were women's swimming competitions first included?

- ☐ 1896
- ☐ 1908
- ☐ 1912
- ☐ 1924

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Appendix C Open-ended Questions

Almost all respondents who were re-surveyed (94%) provided an answer to the open-ended question. We utilized GPT-4 to identify the key themes and analyze their prevalence and content. Four main themes emerge: frequency of discussion, content of discussion, roles in finance, and conflict. We created word clouds separately for men and women for each of these themes in Figure C1.

Regarding how often couples discuss financial issues, both male and female participants reported discussing finances regularly, though the frequency varied. Men typically mentioned monthly discussions, often centered around significant financial decisions such as investments, savings, and retirement planning. Women, on the other hand, tended to report more frequent conversations, focusing not only on long-term goals like retirement but also on managing day-to-day expenses and savings strategies. This difference in frequency is seen in the word clouds where “monthly” is the most common word for men, while “weekly” is the most common for women. In many cases, women handle the finances and keep their partners informed regularly, discussing how to pay bills and save for major life events.

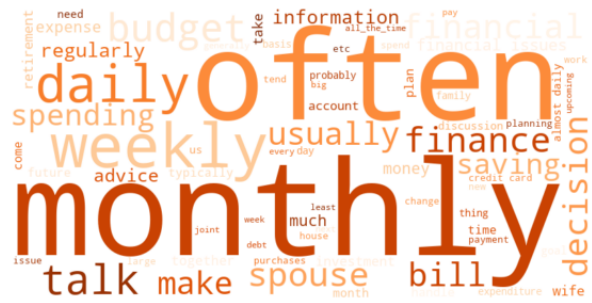
The content of financial discussions for both men and women frequently revolved around managing savings, retirement planning, paying bills, and handling both short-term and long-term financial goals (see words “saving”, “retirement” and “bill” in women and men’s word clouds). Men also highlighted conversations about managing expenses for vacations and deciding where to allocate or save money. Similarly, women emphasized the importance of savings, timely bill payments, and ensuring that both partners stayed on track with financial goals, particularly reducing debt and maintaining financial stability.

The roles in managing household finances varied for both men and women. Some men reported that their spouses primarily handled financial matters, with them checking in periodically or for major decisions. Others noted a shared approach, where both partners contributed equally to decision-making, especially for large purchases or investments. Women, meanwhile, often mentioned taking the lead in financial decisions, managing day-to-day finances, and offering financial advice, which their spouses typically followed. In many cases, financial decision-making was collaborative, with both partners involved in major financial planning.

Both men and women acknowledged that financial discussions could sometimes lead to conflict. Men often mentioned stress arising from differing financial philosophies, with one partner more focused on saving while the other preferred to spend more freely. These differences sometimes caused tension, but couples often worked through it by making joint decisions. Similarly, women reported that financial stress could lead to arguments (see word clouds), with some preferring to avoid financial discussions altogether to prevent conflict. In cases of conflict, both partners generally sought ways to find common ground and shared responsibility for managing finances.

(A) Frequency of discussions

Men



Men



Men



Men



61